

The impact of indoor solid fuel use on the stunting of Indian children

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Abstract

About half of Indian children under the age of three years are stunted. We use data from the 2005-2006 National Family Health Survey in India to assess the link between being exposed to household air pollution from burning solid fuels and child stunting. We tackle the endogeneity issue between child stunting and household fuel choice with the help of an instrumental variables approach. Moreover, we study the potential heterogeneous impacts of household air pollution and other driving factors on stunting in a quantile regression setting. We find strong evidence that exposure to solid fuel smoke increases the probability of being stunted and severely stunted among Indian children, as well as reduces the height-for-age measure for those children. Our results also indicate that the impact is stronger in the central part of the height-for-age distribution, and not so much in the extremes. Given that this is where the vastest number of children are situated, our results point also to the large scale of the required policy intervention. Incentivizing a diminished exposure to household air pollution is expected to be no small undertaking that targets only some pockets of the population, but rather a pan-Indian project.

Keywords: Indoor air pollution; solid fuel; stunting; height-for-age score; instrumental variables; ventilation.

JEL Classification Codes: C35, C36, I15, I18, I32, I38.

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1 Introduction

The reliance on solid fuels for cooking and heating in developing countries has exacerbated the problem of disease among their populations (Fullerton et al., 2008). The combustion of solid fuels releases particulate matter consisting of carbon monoxide, benzene, formaldehyde, and other toxins into the surrounding air (Smith, 2000, p. 13287). The by-product of these toxins are damaging to the health of individuals exposed, with problems compounded by the lack of adequate ventilation in indoor areas. Studies have shown that exposure levels in unventilated households using solid fuels are 10 – 50 times greater than exposure in ventilated or outdoor areas (Smith, 2000, p. 13287). Moreover, household air pollution (HAP) levels significantly depend on the type of fuel used for cooking and heating, with some fuels generating up to ten times more particle matter than others (Smith et al., 2011). Solid fuel types include wood products, agricultural crop residue, and animal dung; cleaner fuels (non-solid) consist of electricity, liquid petroleum gas, and biogas.

Among the studies that analyze the negative health consequences of being exposed to HAP caused by burning fuels, a plethora of them highlights the contribution to acute respiratory problems (Pandey et al., 1989; Armstrong and Campbell, 1991; Mishra and Retherford, 1997; Sharma et al., 1998; Smith, 2000; Ezzati and Kammen, 2001; Brunekreef and Holgate, 2002; Balakrishnan et al., 2002; Duflo et al., 2008; Torres-Duque et al., 2008; Tielsch et al., 2009; Yu, 2011; Prasad et al., 2012; Upadhyay et al., 2015).

This study focuses on India where HAP is a major health concern and ranks third in risk factors for disease, behind high blood pressure and high blood sugar (Forouzanfar et al., 2015). In India, children are particularly vulnerable to the harmful effects of indoor air pollution due to their tendency to stay indoors and be carried by their mothers while cooking (Mishra and Retherford, 2007, p. 377).

In this paper, we focus on a particular health impact of solid fuel use, namely its link with child stunting. Stunted growth is at an alarmingly high level in India; based on the 1998 National Family Health Survey (NFHS-2), about half of the Indian children under three years of age were displaying severely reduced height-for-age ratios (Mishra and Retherford, 2007). More worryingly, despite high economic growth during the last few decades, the propensity for stunting seems to have remained at high levels in India, possibly signaling that (i) India's growth is not inclusive and does not benefit the poorest part of the society where stunting tends to prevail, or that (ii) there are some *sticky* drivers of stunting, such as everyday family cooking and eating practices, that are well rooted in family or community culture and do not change with income variations. In support of the weak link between income growth and stunting rates, Vollmer et al. (2014) show that among the 36 countries in their sample (including India), no significant association can be found between the change in per capita GDP and the change in the percentage of stunted children over the 1990-2011 horizon.

Only a few studies so far have specifically focused on the link between stunting and exposure to indoor air pollution. [Mishra and Retherford \(2007\)](#) use a multinomial logistic regression approach to estimate the impact of solid fuel use on the predominance of anemia and stunting among Indian children. Employing data from the 1998-1999 National Family Health Survey, they find stunting to be significantly more prevalent among children from households using solid fuels. Relying on a similar methodological approach, [Machisa et al. \(2013\)](#) analyze the association between solid fuel use and stunting in children from Swaziland. The authors find no significant evidence of a negative impact of solid smoke on height-to-age ratios after adjusting for child characteristics such as sex, age, birth weight, and preceding birth interval.

Our contribution to the literature is threefold. First, we analyze the most recent data available from the 2005-2006 India's National Family Health Survey (hereafter NFHS-3) for children aged 3 years old and younger, and compare our findings with the related literature that used the 1998-1999 data. While our main variable of interest is the burning of solid fuels inside the house, we account for other possible causes of stunting that include residential, parental, and child-specific features, as well as wealth and nutritional intake. Second, we opt for an instrumental variables estimation approach. As households are likely to take decisions that affect both child health and domestic purchases (such as fuel type for cooking) simultaneously, endogeneity concerns arise. The instrumental variables approach is employed in this setting to help us tackle endogeneity issues, reduce the bias of the estimated coefficients, and identify the causal impact of cooking with solid fuel on the stuntedness of children. We expect our study to deliver higher precision estimates of the impact of solid fuel on stunting than previous literature. The estimation is first focused on the conditional mean of stunting and, second, on the conditional quantiles of stunting. Finally, as opposed to most previous literature, we analyse the height-for-age score in addition to the binary stunted and severely stunted measures. This allows us to check how household characteristics relate linearly to the continuous HAZ measure, and not only their effect on driving HAZ beyond the stunting thresholds.

We find strong evidence that exposure to smoke from burning solid fuel leads to a lower height-for-age ratio and can explain the prevalence of stunting and severe stunting among Indian children. Being exposed to solid fuel smoke increases the probability of being stunted by about 6.5% on average for Indian children below three years old. This effect is half as impactful as following an inappropriate nutrition, which is the usual top concern when trying to fight stunting. Other contributing factors to reduced height-for-age are poor kitchen infrastructure, a higher child birth order, and low levels of mother's education and health status. Moreover, we find that residing in rural areas is associated with reduced stunting, possibly due to lower outdoor air pollution levels than in urban centers, as well as other unobserved influences. The instrumental variables quantile regression results also indicate that the impact of solid fuel is stronger in the central part of the height-for-age distribution,

while less so in the extremes. Given that this is where the vastest number of children are situated, our results point also to the large scale of the required policy intervention. Providing incentives to attain a lower exposure to household air pollution is expected to be no small undertaking that targets only some pockets of the population, but rather a pan-Indian project.

The remaining of this paper is organized as follows. Section 2 introduces the outcome variable and key statistics related to it in the Indian context. Sections 3 and 4 lay the foundations for our theoretical and methodological approach. Section 5 provides details on the dataset used and the key sample statistics. We present the main findings in section 6. Finally, Section 7 concludes and, based on our findings, discusses some potential avenues to lower households' exposure to indoor air pollution and avoid severely reduced growth rates.

2 Measuring stunted growth

In a well-nourished and healthy population, there is a statistically predictable distribution of height for children of a given age. The standard index used for physical growth – height-for-age – reflects the long-term effects of malnutrition, and indicates the skeletal growth of the child. The height-for-age (HAZ) measure is expressed in standard deviation units (Z -score) from the median of a reference population, and is defined in Eq. (1):

$$HAZ_i = \frac{\text{Height}_i - \text{Median}(\text{reference population})}{\text{Standard Deviation}(\text{reference population})}. \quad (1)$$

where HAZ_i is the height-for-age indicator of child i , whose height is given by Height_i . According to WHO standards, a height-for-age Z -score less than -2 indicates being stunted, while a negative Z -score of less than -3 indicates being severely stunted. We rely on the most recent international reference population, which was released by WHO in 2006 (WHO Multicentre Growth Reference Study Group, 2006).¹ The Z -score of the reference population is normally distributed; then, a child in the reference population will have a chance of less than 2.3% to be stunted (Imai et al., 2014).

In our analysis, we consider three alternative definitions of stunting. We rely on two indicator variables (stunted and severely stunted) and the continuous height-for-age Z -score (HAZ). First, using three different measures helps with examining the robustness of our analysis regarding the measurement of stunting. Second, we intend to capture not only the factors that contribute to pushing height-for-age behind a certain threshold, but also to understand the sensitivity of the continuous height-for-age variable to different factors.

¹The new standard is based on children of non-smoking mothers, around the world (Brazil, Ghana, India, Norway, Oman, and the United states) who are raised in healthy environments and are fed with recommended feeding practices (exclusive breastfeeding for the first 6 months and appropriate complementary feeding from 6 to 23 months).

Stunted children tend to have both physical and cognitive developmental delays, including delayed walking, delayed speech development, and reduced school performance. They also experience higher rates of mortality and morbidity, including diabetes and hypertension. Moreover, lower height-for-age levels tend to have suboptimal long-term consequences for an individual’s welfare (Glewwe et al., 2001; Akresh et al., 2011). Hoddinott et al. (2013) find the number of years of education, the scores in aptitude tests, the characteristics of marriage partners, and women’s age at first birth to be on average lower for individuals that had low height-for-age scores as children.

In India, a very high incidence of low height-for-age ratios has been documented, with 46% of children under three years of age being stunted according to the 1998-99 National Family Health Survey (NFHS-2) (Mishra and Retherford, 2007). Our analysis relies on India’s National Family Health Survey (NFHS-3) conducted in 2005-2006.² Figure 1 shows the distribution of the HAZ score in our sample. The incidence of stunting appears to be similar to the 1999 levels, with about half of the Indian children younger than three years old being classified as stunted. Moreover, an alarmingly high percentage (25%) is classified as severely stunted. It is also concerning that the mean value of the HAZ score is very low and drawing close to the stunting threshold.

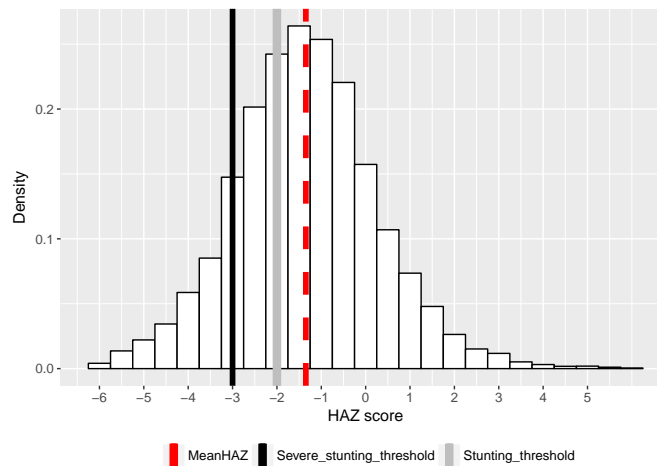


Figure 1: Distribution of HAZ score for children 36 months old and younger in India, based on NHFS-3. The vertical lines mark the sample mean, and the cut-offs for stunting at -2 and severe stunting at -3.

²The NFHS-3 is a nationally representative sample covering all 29 Indian states with an overall response rate of 92.4%. The survey contains information on infant and child mortality, maternal and child health, nutrition, anemia, and the utilization and quality of health and family planning services. NFHS-3 adopted a two-stage sample design in most rural areas while a three-stage sample design was adopted in most urban areas. More details about the survey and its design are provided in IIPS (2007).

3 A theoretical model of child health

Child nutrition status is determined by a complex set of factors (Fenske et al., 2013), some intrinsic to the child (such as age and gender) and some characteristic to the family and community in which the child is raised (such as maternal education and access to safe drinking water). Following previous literature (Becker, 1981; Strauss and Thomas, 1998; Bassolé et al., 2007), we rely on a simple and widely diffused model to capture the demand for child health. In a framework of unitary household models,³ each household maximizes its total utility function by taking decisions regarding children’s health (S),⁴ leisure (L), and the consumption of goods and services (G). The optimization problem can be written as:

$$\max_{S,L,G} U(S, L, G; X_P, \mu) \quad (2)$$

$$\text{such that: } B = p_S \cdot S + p_L \cdot L + p_G \cdot G \quad (3)$$

where U is a continuous utility function, increasing in the inputs (S , L , and G), and quasi-concave. X_P is a vector of parental characteristics, such as mother’s years of education and current employment status. μ denotes the unobserved heterogeneous preferences of the household. Utility is maximized subject to a budget constraint captured in Eq. 3, where B is the total household budget allocated to children’s health, leisure, and the purchase of goods and services. p_S is the price per unit of child health, such as the price of food for feeding the child or the price of medicines in case the child gets sick. The unit prices for leisure and goods and services are p_L and p_G , respectively.

Child health (S) is modeled as a production function that depends on household (X_H), child (X_C), and parental (X_P) characteristics:

$$S = f(X_H, X_C, X_P) \quad (4)$$

Solving the optimization problem of the household (Eq. 2) will lead to an optimal child health level (S^*) given by:

$$S^* = S^*(X_H, X_C, X_P, B, p_S, p_G, \mu) \quad (5)$$

Equation 5 gives insights into the variables that impact child health. For instance, a higher level of education of the parents (reflected in X_P) would lead to a better health (higher height-for-age). Our variable of interest – relying on solid fuels as primary cooking fuel – enters the set of household characteristics (X_H) and we expect that using solid fuels will lead to a worsening in child health, i.e. a reduction in the height-for-age score.

³Unitary household models assume that both parents have identical preferences, resources are pooled, and decisions are taken jointly (Imai et al., 2014).

⁴It is assumed that children are not decisions makers. Instead, parents take decisions that maximize children’s health and nutritional status, given certain constraints.

4 Empirical methodology

In this section, we discuss the econometric models used to estimate the effects of different types of cooking fuel (solid vs. non-solid) on child stuntedness. In a first step, we assume that the impact of being exposed to indoor pollution from solid fuels is constant over the entire height-for-age distribution and estimate the mean impact on stunting. In Section 4.2, we study the impact of solid fuel on different segments of the children population via a quantile regression approach.

In order to estimate Eq. 5, we follow Chay and Greenstone (2003) and assume that the impacts of explanatory variables are linear and additional. The general linear model of child stunting has the following basic specification:

$$S_i = \beta_0 + \beta_1 \text{SolidFuel}_i + \beta_2 X_i + \varepsilon_i, \quad (6)$$

The outcome variable, S_i , captures the long-term nutritional status of child i , as indicated by three measures. First, we rely on the continuous height-for-age score, with $\text{HAZ} \in [-6, 6]$ in our sample. The remaining two measures are constructed as dummy variables and reflect whether or not child i is stunted ($S_i = 1$ if $\text{HAZ} < -2$ and $S_i = 0$ otherwise) or severely stunted ($S_i = 1$ if $\text{HAZ} < -3$ and $S_i = 0$ otherwise).

X_i is a vector of independent variables that determine child stunting, and include household, child, and parental characteristics, i.e. $X_i = \{X_H, X_C, X_P\}$ from Eq. 4 and detailed further in Section 5. State fixed effects are also accounted for in the model in order to capture any differences across states that may arise due to, e.g., state-specific policies or other unobserved state-level factors.

β_1 is the main coefficient of interest and captures the true impact of indoor air pollution from solid fuel use on the stunting measure of child i . SolidFuel_i is a binary variable that takes a unit value if child i resides in a household that uses solid fuel as the main cooking fuel, and zero otherwise. The ordinary least square estimator of β_1 requires that $\mathbb{E}[\text{SolidFuel}_i \cdot \varepsilon_i] = 0$ for a consistent estimation. If, in Eq. 6, there are omitted factors that covary with the solid fuel variable, then the estimated coefficients via least squares will be biased.

The next section explains why endogeneity is a likely problem for our estimation and how we suggest to tackle it via an instrumental variables approach.

4.1 Instrumental variables estimation

In our setting, endogeneity concerns arise from two directions. First, households are likely to take decisions that affect both child health and domestic purchases (such as solid fuel types) simultaneously. One can think about a family's resources (such as time or finances) as being allocated between caring for children and obtaining the fuel needed for cooking. If children are unhealthy, their parents might choose to spend less time outside gathering fuel

(especially solid fuel), and more time inside caring for the children. Tending to sick children could also reduce cooking time. Second, the endogeneity concerns are reinforced by the fact that the fuel choice variable only captures the main fuel used in the household at the time of the survey, and is likely to suffer from random measurement error. As households tend to rely on multiple fuel options, both solid and non-solid, we lack information on what enters the fuel mix and in which proportion beyond the primary component, and whether this is predominantly clean or less so. Overall, the potential existence of both omitted variables and measurement error motivate us to rely on the instrumental variables approach in order to tackle the endogeneity that is likely to bring bias on all model coefficients, not only of the endogenous variable (Wooldridge, 1995).

Besides the theoretical evidence for endogeneity, we assess its presence in our model via a series of tests. First, we run a test of randomization (Stock and Watson, 2003, p. 526), which reveals that the binary variable $SolidFuel_i$ is not random. Second, following Chay and Greenstone (2003), we examine the correlation between fuel choice and predicted outcome variables (stunted indicator, severely stunted indicator, and height-for-age score). We first predict the outcome variable on all exogenous variables in the structural equation (Eq. 6), excluding fuel choice. We then regress the predicted outcome variable values on fuel choice and find a strongly significant relation, signaling once again that OLS is likely to lead to severely biased estimates and reinforcing the choice for the IV approach. The results of the two endogeneity tests are detailed in the Appendix.

The challenge that arises in implementing the instrumental variables approach is to find a strong and valid instrument, which is (i) significantly correlated with the endogenous regressor, after controlling for the exogenous explanatory variables of the structural equation (*relevance criterion*), and (ii) uncorrelated with the error term of the main equation (*exclusion criterion*); see Kennedy (2003). In other words, the instrument should be correlated with solid fuel use, our endogenous variable, to satisfy the relevance condition, but affect child stuntedness only through its impact on solid fuel use in order to satisfy the exclusion criterion.

To find a strong and valid instrument, the drivers of household's fuel choice need to be identified first. The instrument will then be selected among the factors that influence the choice of fuel, but do not impact child stunting directly. Each fuel type has its own private cost, given by the sum of its financial cost (purchase price) and its opportunity cost (the foregone resources, such as own time needed to collect fuel). The financial cost is oftentimes zero for households living in rural areas that rely on solid fuels, as they get to collect it from the commons. However, the opportunity cost in terms of collection labor time might be large; gathering firewood can require up to six hours daily, depending on distance to forest, forest density or degradation state. For other solid fuels, the opportunity cost is related to their alternative uses: dung is a valuable fertilizer, while crop residues can serve as animal fodder. In urban areas, where relying on commercial fuels is more common and non-solid fuels are predominant, opportunity costs tend to be low compared to purchase prices, unless

the latter are lowered through governmental subsidies (Heltberg et al., 2000).

Besides financial and opportunity costs, other factors have been found to play a significant role in the decision for fuel types. Cultural preferences⁵ or a strong (sometimes religion-driven) loyalty to traditional cooking methods can lead households to opt for a fuel type instead of another. On the other hand, using non-solid fuels can be perceived as a symbol of welfare and some households rely on it also to signal their higher social status. Although not a widely spread factor yet, being aware of the negative health consequences of using solid fuels can make households choose cleaner cooking fuels instead of traditional ones. Finally, the choice is further driven by ownership of a stove that runs on cleaner fuels, such as LPG and kerosene. Even when households have a strong preference for using non-solid fuels, the initial costs of purchasing an appropriate stove can be prohibitive, especially in situations of restricted access to credit (Gupta and Köhlin, 2006). It is easy to see then that the private cost of different fuel types varies widely across households. Moreover, each household chooses the type and amount of cooking fuel *'by maximizing its utility subject to a "virtual" or "shadow" price of energy which is unobserved and unknown, except to the household itself, and which varies between households depending on household and village characteristics'* (Heltberg et al., 2000).

*** Modify from here to introduce the chosen instrument(s)***

The latter can be understood in terms of geographical proximity to the resource, such as closeness to the forest (wood), to an agricultural zone (dungcakes and crop residues), or to the market (kerosene and LPG). Fuel availability is likely to influence childhood stunting only indirectly through the impact on fuel choice, making it a potential candidate for a strong and valid instrument. We propose to proxy fuel availability by the share of houses that uses solid fuel as main cooking fuel in a primary sampling unit (PSU), and use it as instrument for the endogenous variable. We define the fraction of houses that uses solid fuel in the PSU (Z) as:

$$Z_{ij} = \frac{\text{Nr. households in PSU } j, \text{ where child } i \text{ resides, that use solid fuel as main fuel}}{\text{Nr. households in PSU } j, \text{ where child } i \text{ resides}} \quad (7)$$

Z_{ij} is constant for all children living in PSU j , i.e. $Z_{ij} = Z_j, \forall i \in [1, N_j]$, where N_j is the total number of children in PSU j .

The proposed instrument satisfies the relevance criterion, being strongly correlated with fuel choice, as illustrated by the results of the first stage estimation; see Section 6. While it is not possible to formally test whether the instrument satisfies the exclusion criterion in our just identified model (with only one instrument for the endogenous variable) (Kennedy, 2003), the condition is fulfilled on theoretical grounds. The prevalence of solid fuel among neighboring households is likely to capture the ease of fuel accessibility in the area (proxim-

⁵Some foods are considered to taste better when cooked with solid fuels than with non-solid fuels, such as chapati in India, Aung et al. (2016).

ity to source), without having a direct impact on childhood stunting.⁶

The IV approach consists in a two-stage estimation of the following model:

$$\text{SolidFuel}_i = \gamma_0 + \gamma_1 Z_i + \gamma_2 X_i + \nu_i \quad (8)$$

$$S_i = \beta_0 + \beta_1 \widehat{\text{SolidFuel}}_i + \beta_2 X_i + \varepsilon_i \quad (9)$$

In the first stage, solid fuel use is regressed on the instrumental variable (Z_i) capturing the fraction of households relying on solid fuel for cooking in PSU_i , and the set of exogenous regressors, X_i , to produce the predicted endogenous variable $\widehat{\text{SolidFuel}}_i$. In the second stage, the dependent variable is regressed on the predicted $\widehat{\text{SolidFuel}}_i$ and the vector of exogenous variables X_i . The IV method aims to isolate a part of the variation in the endogenous explanatory variable SolidFuel_i that is not influenced by the omitted variables, in order to generate consistent estimates. The IV estimation is expected to produce unbiased estimates of the true impact of being exposed to indoor air pollution from burning solid fuels on childhood stunting.

4.2 Quantile regression models

The ordinary least squares and IV models presented above assume a constant association between the explanatory variables and the outcome variable over its entire distribution. The two models will be useful for estimating the mean impact of living in a household that uses primarily solid fuels for cooking on a child's malnutrition status. However, it is likely that children will react differently to the exposure to solid fuels smoke, depending on where they are on the distribution of height-for-age. Such heterogeneous effects of the explanatory variable on the outcome variable can be estimated via a quantile regression (QR) approach, which was first introduced by [Koenker and Bassett Jr \(1978\)](#).

As detailed in [Bassolé et al. \(2007\)](#), the QR approach has a number of advantages compared to the OLS model, as it is less sensitive to outliers, it leads to more robust estimators when the normality assumption is not satisfied, and it performs better when heteroscedasticity is present.

The QR approach has been increasingly applied to understand the impact of different drivers on child stunting in the recent literature. Focusing on child malnutrition in Senegal, [Bassolé et al. \(2007\)](#) finds heterogeneous distributional impacts of access to public infrastruc-

⁶Recent studies, e.g. [Chafe et al. \(2015\)](#), point to the significant contribution of burning solid fuels inside to outdoor PM2.5 levels. This could raise doubts that the chosen instrument satisfies the exclusion criterion, as ambient air quality can also impact child health and stunting rates. However, it has been estimated that burning solid fuels inside contributes up to 10% to global ambient pollution, while the contribution to local air quality depends on various confounding factors (such as topography and weather) and no conclusions have been reached so far on the magnitude of these impacts. With this background, we assume that there are no systematic spillover effects of indoor pollution to the average outside pollution level in the primary sampling unit and maintain our position that the chosen instrument satisfies the exclusion criterion.

ture (such as safe water and health facilities) on child’s stunted growth. [Burchi \(2010\)](#) relies on a QR approach to study the impact on child stunting of the mother’s education level and nutrition knowledge in Mozambique. The study shows that, while mothers’ nutrition knowledge reduces stunting among children at the lower end of the height-for-age distribution, her formal education is more important for increasing height-for-age in the higher quantiles. Similar results related to parents’ education and its impact on child stunting are reported by [Borooah \(2005\)](#) in India, using data collected from 15,000 households in 1994. Relying on the NHFS-3 of 2005-2006 in India, [Kandpal et al. \(2009\)](#) show that maternal health and education are better able to reduce stunting at the lower end of the distribution than on the upper end. Another study employing the NHFS-3 dataset is [Fenske et al. \(2013\)](#), who study the association between various socio-economic factors and childhood stunting in the lower quantiles (15% and 35%) of the height-for-age distribution of 0-24 month old Indian Children.

The overarching message of all QR studies of child stunting is that policy interventions should account for the differential impacts that driving factors may have on child nutrition at different points of the height-for-age distribution. As [Aturupane et al. \(2011\)](#) explain, some interventions might have very limited impacts for reducing the prevalence of stunting on average, but could be highly impactful for children at the lower end of the height-for-age distribution. With this motivation, we will employ the QR approach in our analysis to first assess the potential differential impact of being exposed to solid fuel smoke on child stunting, and second formulate policy recommendations that account for this heterogeneous relation.

5 Data

This section describes the variables employed in the empirical analysis. Key summary statistics at the child level are captured in [Table 1](#). Additionally, we divide the sample according to two criteria: (i) the outcome variable, i.e. child’s stunting status as measured by an HAZ score below -2, and (ii) the endogenous variable, i.e. the main fuel type used in the household for cooking (solid vs. non-solid fuels). Accounting for the two criteria, [Table 1](#) includes subsample summary statistics in order to highlight important differences between the subgroups.

First, [Table 1](#) captures entire sample and subsample statistics for three outcome measures that proxy child health: two binary variables capturing whether the child is stunted and severely stunted, and the continuous height-for-age measure (the HAZ score). In our sample, 38% of the children are stunted while 16% are severely stunted. The average HAZ score is 1.5 points below zero. When distinguishing between children exposed and not exposed to solid fuel smoke, we observe a significantly higher prevalence of stunting and severe stunting in

the former subsample compared to the latter. This observation brings first-hand support to the hypothesis that there is a link between the exposure to solid fuel smoke and adverse health impacts. Our analysis will try to clarify whether this link is causal or the solid fuel use indicator is a proxy for other factors (such as general wealth) that can influence child health.

Second, Table 1 illustrates statistics of other variables that could potentially influence childhood stunting. The main focus of this study is on the link between stunting and household air pollution. HAP levels vary significantly according to the type of fuel used for cooking and heating, with some fuels generating up to ten times more particle matter than others (Smith et al., 2011). The NFHS-3 includes information on households’ primary fuel type used for cooking.⁷ Sorted by their pollution potential from highest to lowest, we divide the fuel types into two broad categories: (i) **solid fuel**, including animal dung, agricultural crop waste, straw/shrubs/grass, wood, charcoal, and coal/lignite, and (ii) **non-solid fuel**, including kerosene, biogas, LPG/natural gas, and electricity.⁸ Our dataset does not provide information about the intensity of child exposure to solid fuels (as measured in hours of daily exposure, for example). Thus, we proxy exposure by an indicator of fuel type: solid versus non-solid. In support of this, Ramaswamy et al. (2004) find that the solid fuel indicator is a good predictor of high concentrations of particulate matter in the living space.

Figure 2 illustrates the distribution of the HAZ score by solid fuel use in our sample. While this graph does not imply any causation, it is quite striking to see the association of solid fuel use with the HAZ score, as marked by the different average HAZ scores in the two subsamples.

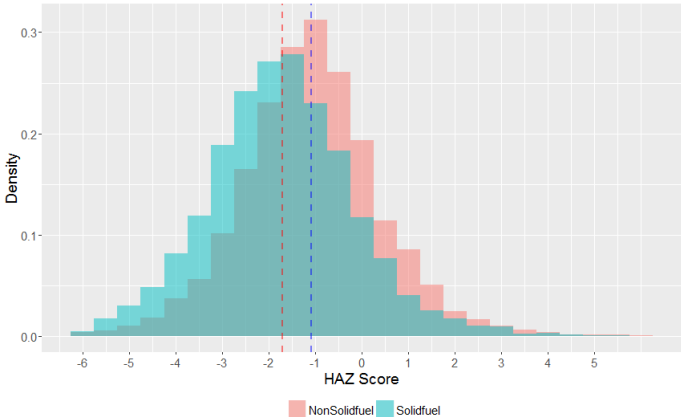


Figure 2: Distribution of HAZ score by main fuel type used in the household.

Considering the exposure to HAP as a potential factor for stunting, one also needs to account for the cooking practices of the household which, directly or indirectly, may reduce the exposure and impact of IAP. In particular, we account for possible measures of indoor

⁷The survey asked each respondent “What type of fuel does your household mainly use for cooking?”.

⁸The residual category of other fuels was excluded from the analysis.

ventilation and the configuration of the cooking space. For this, we use binary variables that denote whether the house has any windows and if the kitchen is separated from the remaining living space. Finally, to control for other possible sources of indoor air pollution, we include an indicator variable to signal whether the child is exposed to tobacco smoke in the household.

Beyond the impact of household air pollution, our analysis attempts to account for other variables that could potentially impact stunting. Based on previous contributions to the literature, we group the controls into (A) residential, (B) childhood, and (C) parental characteristics. As Table 1 indicates, the difference in means between the subgroups is significantly different from zero, for all considered variables, with very few exceptions. This indicates that indeed the characteristics of the stunted are different than those of non-stunted children; the same holds true regarding the differences between the group of children exposed to solid fuels smoke and those not exposed to it.

A. Residential characteristics

Previous literature finds that stunting tends to be significantly associated with various household characteristics, such as family income, household crowding, house construction type, access to safe water, geographical region, and setting (urban or rural) (Shah et al., 2003; Machisa et al., 2013; Fenske et al., 2011). As NFHS-3 lacks information on actual family income, we proxy it by including several indicator variables capturing the presence of durable assets (such as refrigerator, radio, water pump, etc.).

Table 1 shows that, in our sample, stunted children tend to live in households where (i) solid fuels are the primary fuel type used, (ii) the cooking space is not separate from the remaining living space in the house, (iii) the house is less likely to have any windows, (iv) the house is located in a rural area, (v) the floor of the house is more likely to be made of clay, soil, sand, or dung, (vi) household members are more likely to practice open defecation, and (vii) the family owns in general less durable goods.

B. Child's characteristics

Previous studies find age, gender, birth order, birth weight, and recent sicknesses to be significantly correlated with stunting (Jayachandran and Pande, 2015; Machisa et al., 2013; Fenske et al., 2011; Mishra and Retherford, 2007; Adair and Guilkey, 1997). We control for these characteristics and, additionally, include information related to the children's recent health history: whether they had fever and/or diarrhea in the two weeks prior to the survey.

The duration a child is breastfed is another factor found to have influence on the height-for-age score. However, the sign of the relation is the subject of academic debate; see the discussion in Section 6. The WHO recommends exclusive breastfeeding for the first six months of life. After that, children should receive nutritionally adequate and safe comple-

mentary foods, while breastfeeding should continue for up to two years of age or beyond.⁹ We include an indicator that marks whether or not the child has been breastfed for a period of time in line with the WHO recommendations.

Table 1 illustrates that our sample is gender-balanced, and includes children with ages 35 months and younger. In our sample, stunting appears to be more prevalent among older infants and with a daily nutrition based on less varied food. WHO, in its guiding principles for complementary feeding of children, recommends that a breastfed child aged 6-23 months should be fed from 4 or more different food groups. We indicate the diversity of a child's diet by a binary variable, such that a value of 1 is assigned if a child consumed 4 or more food groups in the last 24 hours, otherwise the indicator is zero.¹⁰

C. Parents' characteristics

An important role for the occurrence of stunting appears to be played by some characteristics of the parents (Shah et al., 2003; Rayhan and Khan, 2006; Semba et al., 2008). Our analysis accounts for several aspects related to the mother, such as: body mass index, height, an indicator signaling whether she is currently employed, and the number of years of education.

In our sample, it is striking that mothers have on average only primary education and less than a third are employed at the time of the survey. Distinguishing between stunted and non-stunted children, Table 1 identifies the potential influence on stunting of the number of years of mother's education, whether or not she supplemented her diet with iron during pregnancy, and her current smoker status.

Summing up, analyzing the key statistics in our sample, many socio-economic differences arise between the children exposed to solid fuel and those not exposed to it. Our analysis will control for these discrepancies and attempt to quantify the net impact of being exposed to solid fuel smoke on child stunting.

⁹The recommendation does vary for the developed world. In the US, the American Academy of Pediatrics (AAP) recommends that a child be exclusively breastfed for the first 6 months and should get complementary food thereafter while supporting breastfeeding for the first year.

¹⁰In line with WHO recommendations, the food group categories considered are: (i) milk, cheese, yogurt or other milk products; (ii) bread, noodles or other food made from grains; (iii) oil, fats, butter, or products made from them; (iv) mangoes, papaya, or other Vitamin A rich fruits; (v) beans, legumes/lentils and nuts; (vi) other fruits and vegetables (including green leafy vegetables); and (vii) eggs, meat, poultry, fish, and shellfish (and organ meats).

Table 1: Summary statistics of outcome and explanatory variables

	Entire sample		Stunting status				Fuel type			
	(N = 14,659)		(N = 4,963)		(N = 9,696)		(N = 9,161)		(N = 5,498)	
	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD
I. Outcome variables										
Stunted indicator ($1_{HAZ < -2}$) ^{1,2}	0.38	0.48	1	0	0	0	0.42	0.49	0.25	0.43
Severely stunted indicator ($1_{HAZ < -3}$) ^{1,2}	0.16	0.37	0.43	0.49	0.00	0.00	0.19	0.39	0.09	0.20
Height-for-age score (HAZ) ^{1,2}	-1.50	1.62	-3.08	0.87	-0.54	1.16	-1.65	1.62	-1.04	1.53
II. Explanatory variables										
A. Household characteristics (X_H)										
Household is using primarily solid fuel ^{1,2}	0.74	0.44	0.83	0.38	0.69	0.46	1	0	0	0
Fraction of solid fuel use in PSU ^{1,2}	0.77	0.34	0.83	0.30	0.73	0.36	0.93	0.15	0.30	0.31
Kitchen is separate ^{1,2}	0.52	0.50	0.43	0.49	0.57	0.49	0.44	0.50	0.75	0.43
House has any windows ^{1,2}	0.61	0.49	0.52	0.50	0.67	0.47	0.51	0.50	0.91	0.29
Rural residence ^{1,2}	0.71	0.45	0.77	0.42	0.68	0.47	0.88	0.33	0.23	0.42
Nr. household members ²	6.82	3.25	6.80	3.21	6.83	3.27	7.05	3.28	6.16	3.06
Floor not covered ^{1,2}	0.53	0.50	0.63	0.48	0.47	0.50	0.70	0.46	0.05	0.22
Open defecation ^{1,2}	0.59	0.49	0.69	0.46	0.53	0.50	0.77	0.42	0.10	0.29
Indicator that house has bank account ^{1,2}	0.11	0.31	0.08	0.27	0.13	0.33	0.06	0.24	0.24	0.43
Electricity ^{1,2}	0.63	0.48	0.54	0.50	0.68	0.47	0.50	0.50	0.98	0.14
Radio ^{1,2}	0.31	0.46	0.26	0.40	0.34	0.47	0.27	0.44	0.43	0.49
TV ^{1,2}	0.43	0.50	0.34	0.47	0.49	0.50	0.28	0.40	0.85	0.36
Telephone ^{1,2}	0.12	0.33	0.07	0.26	0.16	0.36	0.05	0.23	0.33	0.47
Computer ^{1,2}	0.02	0.13	0.01	0.08	0.03	0.16	0.00	0.04	0.07	0.25
Fridge ^{1,2}	0.14	0.35	0.08	0.27	0.18	0.39	0.04	0.21	0.43	0.49
Bicycle ^{1,2}	0.54	0.50	0.52	0.50	0.55	0.50	0.55	0.50	0.49	0.50
Motorcycle ^{1,2}	0.20	0.40	0.13	0.33	0.24	0.43	0.11	0.31	0.46	0.50
Car ^{1,2}	0.03	0.16	0.01	0.11	0.04	0.19	0.01	0.09	0.08	0.27
Thresher ^{1,2}	0.02	0.15	0.02	0.12	0.03	0.17	0.03	0.17	0.01	0.11
Tractor ^{1,2}	0.03	0.16	0.01	0.12	0.03	0.18	0.03	0.17	0.02	0.13
Mattress ^{1,2}	0.61	0.49	0.54	0.50	0.65	0.48	0.52	0.50	0.87	0.34
Bed ²	0.87	0.34	0.86	0.35	0.87	0.33	0.85	0.35	0.91	0.28
Chair ^{1,2}	0.52	0.50	0.42	0.49	0.58	0.49	0.40	0.49	0.86	0.35
Table ^{1,2}	0.41	0.49	0.31	0.46	0.47	0.50	0.29	0.45	0.76	0.43
Pressure cooker ^{1,2}	0.38	0.49	0.29	0.45	0.44	0.50	0.21	0.41	0.88	0.33
Electric fan ^{1,2}	0.50	0.50	0.40	0.49	0.56	0.50	0.35	0.48	0.93	0.26
Sewing machine ^{1,2}	0.20	0.40	0.14	0.35	0.23	0.42	0.13	0.34	0.39	0.49
Water pump ^{1,2}	0.12	0.32	0.09	0.28	0.13	0.34	0.10	0.30	0.15	0.36
B. Child's characteristics (X_C)										
Child's age (months) ^{1,2}	16.99	9.66	19.76	8.32	15.33	10.03	16.69	9.69	17.87	9.52
Male child ²	0.54	0.50	0.53	0.50	0.54	0.50	0.53	0.50	0.56	0.50
Birth order ^{1,2}	2.80	1.90	3.07	2.02	2.63	1.81	3.08	2.04	1.99	1.12
Varied food ^{1,2}	0.37	0.48	0.22	0.41	0.46	0.50	0.35	0.48	0.40	0.49
Breastfeeding according to WHO ²	0.84	0.37	0.84	0.37	0.84	0.37	0.89	0.32	0.70	0.46
Child had fever in the past two weeks ^{1,2}	0.18	0.38	0.19	0.39	0.17	0.38	0.19	0.39	0.15	0.36
Child had diarrhea in the past two weeks	0.14	0.35	0.15	0.35	0.14	0.34	0.14	0.35	0.13	0.33
C. Parents' characteristics (X_P)										
Mother's BMI ^{1,2}	19.87	3.24	19.29	2.86	20.21	3.41	19.25	2.63	21.64	4.08
Mother's height ^{1,2}	151.75	5.73	150.49	5.56	152.51	5.69	151.31	5.65	153.02	5.76
Nr. years of mother's education ^{1,2}	4.68	5.06	3.36	4.37	5.47	5.28	3.08	4.12	9.26	4.71
Mother is currently working ^{1,2}	0.28	0.45	0.33	0.47	0.25	0.43	0.32	0.47	0.15	0.35
Mother took Iron during pregnancy ^{1,2}	0.66	0.47	0.61	0.49	0.70	0.46	0.60	0.49	0.84	0.37
Mother is smoking ^{1,2}	0.10	0.30	0.12	0.33	0.09	0.28	0.12	0.33	0.04	0.19

Notes: Reported statistics are representative for the sample used in the regressions models. The initial sample size of NFHS-3 is of 51,555 children; however, due to missing observations of many control variables, the final sample size used in the empirical analysis is heavily restricted. The total sample size of 14,659 observations contains complete information on all variables used in the analysis. Coefficients reflect survey household weights. The superscripts (1, 2) placed next to each variable mark whether the difference in means between subcategories is statistically significantly different from zero, such that a subscript of 1 indicates that, for that specific variable, the difference between the mean of the stunted group and the mean of the not-stunted group is significantly different from zero; a subscript of 2 indicates that the difference between the mean of the group exposed to solid fuel smoke and the mean of the group not exposed to solid fuel smoke is significantly different from zero.

6 Results

6.1 Mean impact of solid fuel smoke exposure on stunting

This section presents the empirical results from the estimation of the long-term nutrition status of children, as measured by three outcome variables: an indicator of stunting, an indicator of severe stunting, and the continuous HAZ score. For each of the three dependent variables, we present the results from the linear estimation and the instrumental variable approach. Table 2 presents the results of the six models. Below we provide an overview of the factors that contribute to stunted growth, with an emphasis on the relationship between stunting and the exposure to smoke from burning solid fuels. The robustness of the linear probability models is strengthened by the results from logistic regression models of stunting and severe stunting; see Table A.3 in the Appendix.

The results of the first-stage instrumental variables regression are presented in Table 3. As expected, the fraction of households using solid fuel in the same PSU has a strong and highly significant effect on the solid fuel use by a household in the same PSU. The value of the F -statistic is far above the rule-of-thumb value of 10 used as standard for models with one instrument for one endogenous variable. The test points to a strong instrument, indicating that the IV results are not affected by weak-instrument bias.

Solid fuel as primary fuel type used in the household

Our main focus is on the link between a child's nutrition status and the main fuel type used in the household where the child resides. As the instrumental variable models reveal, we find evidence that being exposed to solid fuel smoke significantly reduces the HAZ score and increases the probability of being severely stunted.

Taking into account again the sample statistics in Table 1, we see that the difference in means between the HAZ score of children exposed to solid and those not exposed to it is 0.61.¹¹ Moreover, the OLS model estimates that up to 0.13 score points (or 21.31%) of this difference is due to exposure to solid fuel, after accounting for all other driving factors. Employing the estimate from the IV model, about 0.265 score points or almost 43.44% of the difference between the HAZ scores is due to exposure to smoke from burning solid fuels.

Our results reinforce the findings of Mishra and Retherford (2007) that high indoor air pollution levels can contribute to chronic nutritional deficiencies in young children. While their conclusions were only based on correlations, the instrumental variables approach employed in our paper can finally ascertain the existence of a causal link. Moreover, as their study was based on data collected 7-8 years prior to ours, comparing the two results also highlights that, unfortunately, households' reliance on solid fuel continues to negatively af-

¹¹The mean HAZ score is -1.65 for children exposed to solid fuel smoke, while the mean HAZ score is -1.04 for children not exposed to it.

fect children's health and contributes largely to the prevalence of stunting. Over the last two decades, the increase in international awareness regarding the negative impacts of household air pollution on health seems to not have been strong enough to motivate parents to keep their children unexposed to indoor pollution.

The residential characteristics marking indoor ventilation possibilities are essential for understanding the influence of solid fuel exposure on child stunting. We find that the availability of windows in the house and having the kitchen in a separate room than the rest of the living space significantly reduce the prevalence of stunting and severe stunting, and increase the HAZ score. It remains, however, difficult to distinguish if these house configurations indicate ventilation opportunities or if they merely constitute proxies of general household health.

Other residential characteristics

Our results show that several residential characteristics play an important role for the prevalence of stunted growth. We reach three conclusions. First, residing in rural areas seems to be associated with reduced stunting, after accounting for other socio-environmental factors, possibly due to lower outside air pollution levels and lower population densities in villages than in cities, and potentially other unobserved influences. With many studies focusing on reducing indoor air pollution in rural households ([Balakrishnan et al., 2004](#); [Parikh et al., 2001](#)), our study points to the importance of addressing solid fuel burning in urban settings as well, where the impacts are augmented due to higher outdoor air pollution levels.

Second, we control for other potential sources of health problems. For children living in households that practice open defecation, the incidence of stunting is higher, but the association is not statistically significant. This finding is in contrast to recent studies (e.g. [Spears et al. \(2013\)](#)) that underline the negative consequences of open defecation on stunting.

Third, there appears to be a reduced prevalence of stunting in wealthier households, as indicated by the coefficients of the durable goods (see [Table A.2](#) in the Appendix).

Child's characteristics

Controlling for childhood characteristics, our analysis reinforces previous findings that stunting is more prevalent among older children, who are higher in the birth order, who suffered loose bowels episodes in the recent period, and are undergoing a poorly diversified diet.

Being higher in the birth order or having a twin can be associated with a tighter allocation of available resources as families get bigger and a higher prevalence of stunting among later born children. Moreover, cultural factors can again play a role, with Indian families tending to show preference for first borns (see [Jayachandran and Pande \(2015\)](#)).

Finally, we find a negative relation between the duration of breastfeeding and child stunting. This health puzzle has been the subject of extended research that points to the existence

of reverse causality (Marquis et al., 1997; Simondon and Simondon, 1998) or omitted variable bias (Atsbeha et al., 2015) when modeling the breastfeeding – stunting relation.¹²

Parents' characteristics

The parental features that significantly reduce the prevalence of stunting are found to be related to parents' health, as captured by the mother's body mass index and her height, and socially-environmentally driven, such as mother's years of education. Moreover, children of mothers who have taken iron supplements during pregnancy are less likely to be severely stunted. The exposure to tobacco smoke does not appear to have statistical power in explaining stunting for our sample.

¹²Marquis et al. (1997) and Simondon and Simondon (1998) find that mothers tend to prolong breastfeeding for children that have poor growth and health. Atsbeha et al. (2015) control for the nutritional adequacy of complementary foods and find no evidence of adverse growth effects from prolonged breastfeeding.

Table 2: Regression models of stunting and severe stunting

	LPM		OLS	IV		
	$1_{\text{HAZ} < -2}$ (\mathcal{M}_1)	$1_{\text{HAZ} < -3}$ (\mathcal{M}_2)	HAZ (\mathcal{M}_3)	$1_{\text{HAZ} < -2}$ (\mathcal{M}_4)	$1_{\text{HAZ} < -3}$ (\mathcal{M}_5)	HAZ (\mathcal{M}_6)
Intercept	2.277*** (0.147)	1.236*** (0.117)	-8.694*** (0.475)	2.266*** (0.148)	1.215*** (0.117)	-8.623*** (0.478)
A. Household characteristics (X_H)						
Household is using primarily solid fuel	0.045*** (0.017)	0.017 (0.013)	-0.130** (0.054)	0.065* (0.034)	0.058** (0.026)	-0.265** (0.111)
Kitchen is separate	-0.042*** (0.012)	-0.016* (0.009)	0.088** (0.037)	-0.041*** (0.012)	-0.015* (0.009)	0.087** (0.037)
House has any windows	-0.034** (0.013)	-0.030*** (0.011)	0.135*** (0.042)	-0.034** (0.013)	-0.030*** (0.011)	0.134*** (0.042)
Rural area	-0.021 (0.014)	-0.018* (0.011)	0.075 (0.047)	-0.026 (0.016)	-0.029** (0.013)	0.110** (0.053)
Nr. household members	0.003* (0.002)	0.000 (0.001)	-0.010* (0.006)	0.003 (0.002)	-0.000 (0.001)	-0.009 (0.006)
Floor not covered	0.005 (0.015)	0.005 (0.012)	0.020 (0.049)	0.003 (0.016)	0.002 (0.012)	0.031 (0.049)
Open defecation	0.017 (0.016)	0.010 (0.012)	-0.027 (0.049)	0.014 (0.016)	0.004 (0.012)	-0.009 (0.050)
B. Child's characteristics (X_C)						
Child's age	0.009*** (0.001)	0.006*** (0.000)	-0.042*** (0.002)	0.009*** (0.001)	0.006*** (0.000)	-0.042*** (0.002)
Male child	-0.020** (0.010)	-0.013* (0.008)	0.050 (0.031)	-0.019** (0.010)	-0.013* (0.008)	0.049 (0.031)
Birth order	0.002 (0.003)	0.006** (0.003)	-0.001 (0.010)	0.002 (0.003)	0.006** (0.003)	-0.002 (0.010)
Diverse diet	-0.137*** (0.011)	-0.062*** (0.008)	0.466*** (0.036)	-0.137*** (0.011)	-0.063*** (0.008)	0.467*** (0.036)
Breastfeeding according to WHO	0.028* (0.015)	0.039*** (0.012)	-0.197*** (0.051)	0.028* (0.015)	0.038*** (0.012)	-0.193*** (0.051)
Had fever in the past two weeks	0.008 (0.013)	-0.004 (0.011)	-0.065 (0.041)	0.008 (0.013)	-0.005 (0.011)	-0.063 (0.041)
Had diarrhea in the past two weeks	0.032** (0.014)	0.013 (0.011)	-0.152*** (0.043)	0.032** (0.014)	0.013 (0.011)	-0.153*** (0.043)
C. Parents' characteristics (X_P)						
Mother's BMI	-0.009*** (0.002)	-0.004*** (0.001)	0.039*** (0.005)	-0.009*** (0.002)	-0.004*** (0.001)	0.038*** (0.005)
Mother's height	-0.012*** (0.001)	-0.007*** (0.001)	0.047*** (0.003)	-0.012*** (0.001)	-0.007*** (0.001)	0.047*** (0.003)
Nr. years of mother's education	-0.006*** (0.001)	-0.004*** (0.001)	0.017*** (0.005)	-0.005*** (0.001)	-0.004*** (0.001)	0.015*** (0.005)
Mother is currently working	0.014 (0.012)	0.005 (0.010)	-0.040 (0.039)	0.014 (0.012)	0.004 (0.010)	-0.040 (0.038)
Mother took iron during pregnancy	-0.007 (0.012)	-0.031*** (0.010)	0.050 (0.039)	-0.007 (0.012)	-0.031*** (0.010)	0.050 (0.039)
Mother is smoking	0.009 (0.018)	-0.008 (0.015)	-0.064 (0.056)	0.009 (0.018)	-0.008 (0.015)	-0.065 (0.056)
Household assets	Yes	Yes	Yes	Yes	Yes	Yes
State fixed effects	Yes	Yes	Yes	Yes	Yes	Yes
Nr. observations	14,659	14,659	14,659	14,659	14,659	14,659
R ²	0.157	0.106	0.207	0.157	0.105	0.207
Adj. R ²	0.153	0.101	0.203	0.153	0.101	0.203

Notes: The table capture estimation results from linear models of stunting (\mathcal{M}_1) and severe stunting (\mathcal{M}_2), OLS model for the HAZ score (\mathcal{M}_3), and the IV models for each output measure ($\mathcal{M}_4 - \mathcal{M}_6$). Errors were clustered by household, and robust standard errors are presented in parentheses below each coefficient. Significance levels are indicated by ***, **, *, marking the 1%, 5% and 10% levels, respectively.

Table 3: First stage regression of solid fuel use for the instrumental variables approach.

	HAZ (\mathcal{M}_7)
Intercept	0.23*** (0.07)
Fraction of solid fuel in PSU	0.71*** (0.02)
A. Household characteristics (X_H)	
Kitchen is separate	-0.01 (0.01)
House has any windows	-0.00 (0.01)
Rural area	0.00 (0.01)
Nr. household members	0.01*** (0.00)
Floor not covered	0.05*** (0.01)
Open defecation	0.07*** (0.01)
B. Child's characteristics (X_C)	
Child's age	0.00 (0.00)
Male child	-0.00 (0.00)
Birth order	-0.01*** (0.00)
Varied food	0.01 (0.01)
Breastfeeding	0.02** (0.01)
Had fever in the past two weeks	0.01 (0.01)
Had diarrhea in the past two weeks	-0.01* (0.01)
C. Parental characteristics (X_P)	
Mother's BMI	-0.00* (0.00)
Mother's height	0.00 (0.00)
Nr. years of mother's education	-0.01*** (0.00)
Mother is currently working	0.00 (0.01)
Mother took iron during pregnancy	0.00 (0.01)
Mother is smoking	-0.00 (0.01)
Household assets	Yes
State fixed effects	Yes
Nr. observations	14,659
First stage F -statistic	1,641.11
Prob > F	0.0000

Notes: Errors were clustered by household, and robust standard errors are presented in parentheses below each coefficient. Significance levels are indicated by ***, **, *, marking the 1%, 5% and 10% levels, respectively.

6.2 Heterogeneity in impacts of solid fuel smoke exposure on stunting

This section explores the heterogeneity of child stunting determinants on the height-for-age score. Analyzing the differential impact that household air pollution may have on stunting at different points of the height distribution would allow us to draw conclusions regarding which children should be targeted first by policy interventions for stronger outcomes.

The unconditional distribution of the HAZ score is illustrated in Fig. 3; the vertical red lines mark the values of HAZ at different quantiles of interest. The instrumental variable quantile regression results are captured in Table 4; the coefficients of some selected variables of interest are illustrated graphically in Fig. 4. We find that our variable of prime interest, i.e. whether or not the household uses solid fuel as primary fuel for cooking, has significantly distinct impacts on the height-for-age variable at different quantiles of the distribution. Being exposed to solid fuel smoke appears to significantly reduce height-for-age mostly for the middle of the distribution, i.e. for the 0.5, 0.75, and 0.90 quantiles, while the impact remains not significant at the distribution extremes.¹³

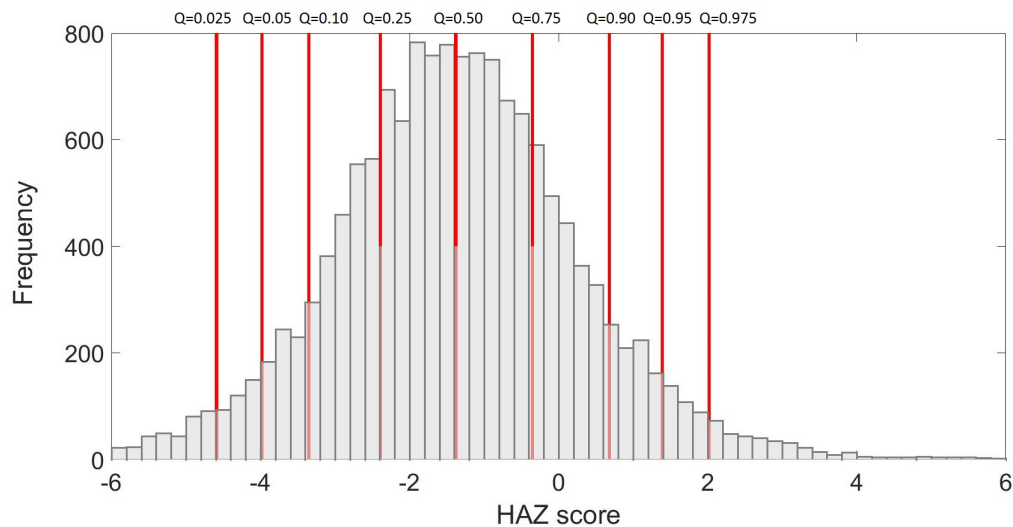


Figure 3: Histogram of the sample HAZ score with markup at different distribution quantiles.

On the left-hand-side of the HAZ distribution, for quantiles of 0.25 and lower, all children are stunted or severely stunted, with an HAZ score below -2 (Fig. 4). It is likely that their health is so weakened by other driving factors that being exposed to solid fuel has no significant additional impact. An alternative hypothesis is that such malnourished children are too weak to spend long hours next to their mothers close to the cooking place and are, thus, less exposed to solid fuel smoke. On the right-hand-side of the distribution, for the 0.95 and 0.975

¹³The robustness of these results is confirmed also the OLS quantile regressions models. Although coefficient signs and significance tends to match between simple QR and IVRQ models, coefficient magnitudes differ, especially for the endogenous variable *Solid Fuel*. Results are captured in Table A.6 in the Appendix.

quantiles, where the HAZ score is 1.39 and higher, we again observe no significant impact of solid fuel smoke on the HAZ score. It appears that the height-for-age score is immune to solid fuel smoke exposure among the healthiest children. Overall, the IVQR analysis reveals that policy interventions regarding the switch from solid to cleaner fuels should be directed neither at the weakest nor the healthiest Indian children. In contrast, most benefits from reducing the use of solid fuels would accrue to children in the middle of the height-for-age distribution. Given that it is here where the vastest number of children are situated, our results also point to the large scale of the required intervention. Incentivizing the switch to cleaner fuels is expected to be no small undertaking that targets only some pockets of the population, but rather a pan-Indian project.

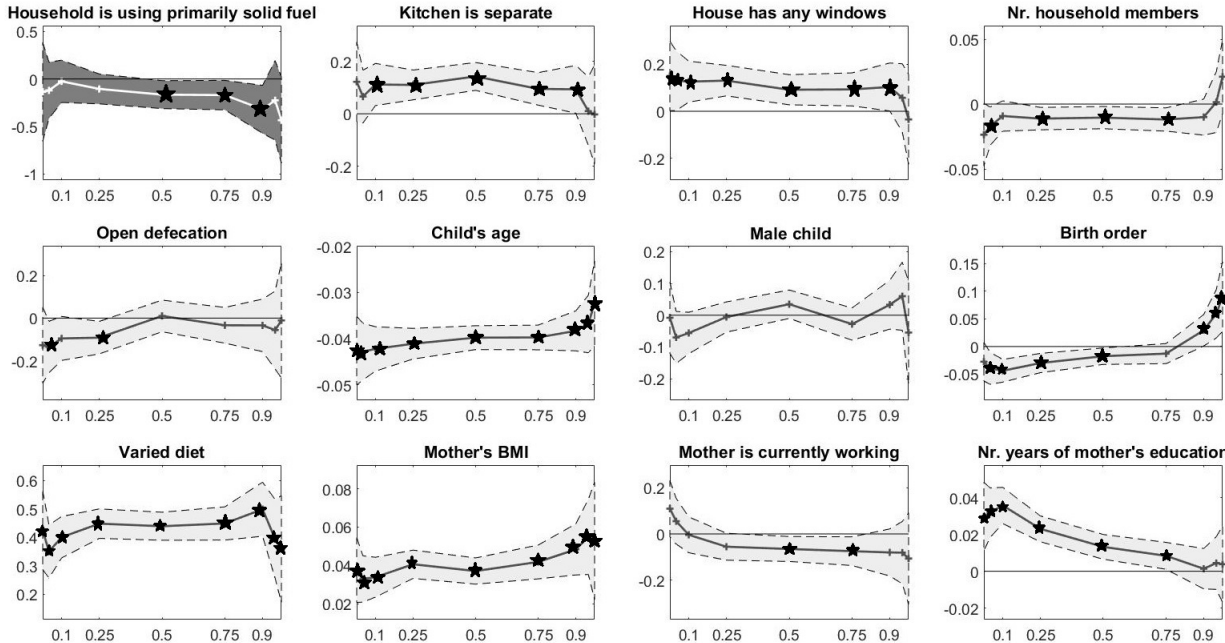


Figure 4: Instrumental variable quantile regressions estimates of selected explanatory variables for child's height-for-age score.

Notes: The figure illustrates the estimated coefficients for selected explanatory variables of the height-for-age score. Results come from fitting IVQR models for the {0.025, 0.05, 0.10, 0.25, 0.50, 0.75, 0.90, 0.95, 0.975} quantiles. 90% confidence intervals are depicted in gray. Statistically significant coefficient are marked with black stars in the graph.

The IVQR analysis allows us to study the association of the height-for-age score with other factors at different points of the HAZ distribution. The results for some selected variables of interest are illustrated in Fig. 4. We draw three conclusions. First, a determinant that could potentially mitigate the impact of household air pollution on stunting is given by ventilation options, such as cooking in a separate room than the rest of the living space and having any windows in the house. Similarly to solid fuel, these two avenues for better venti-

lation have no significant association with the HAZ score at the extremes of the distribution, but do so for the more central quantiles.

Second, some of the variables that had no significant impact on HAZ in the simple IV models, turn out to play a significant role for some parts of the distribution, among which the number of household members, the open defecation indicator, child's birth order, the breastfeeding indicator, the number of years of mother's education, and the indicator of whether the mother is currently working. In particular, breastfeeding appears to have a beneficial impact on children at the bottom of the height-for-age distribution, but detrimental for the higher quantiles; see again the discussion on the literature debate regarding breastfeeding in Section 6.1. Although highly significant in the simple IV regressions models, the IVQR regression analysis reveals that the number of years of mother's education is positively associated with child's nutritional outcomes only for quantiles of 0.75 and below. In contrast, it appears that, for the higher end of the HAZ distribution, maternal education plays no significant role.

Third, some representative variables that had a strongly significant association with height-for-age in the simple IV models continue to exhibit the same strong correlation over the entire HAZ distribution, as highlighted by the IVQR models. In particular, child's age, having a diverse diet, and mother's BMI and height covary positively with height-for-age at all considered quantiles.

Overall, the instrumental variables quantile regressions offer a more complete picture of the relation between child's height-for-age and the explanatory variables, and at times point to the heterogeneity of this relation along the HAZ distribution. Any policy interventions trying to correct the high prevalence of stunting among Indian children should take into account the differential impacts of the driving factors. In particular, we have seen that the switch towards cleaner fuels is expected to bring the highest benefits if incentives programs are directed towards the middle of the height-for-age distribution, where the sensitivity of HAZ appears to be high to the choice of the household fuel. Our results contrast to some extent the findings of [Fenske et al. \(2013\)](#), who employ additive quantile regression models and study the link between socio-economic factors and childhood stunting at the lower quantiles of the height-for-age distribution. The authors find that the association of most variables with stunting on lower quantiles is similar to their impact on the population mean. We believe the difference is driven by the empirical approach choice, additive QR versus IVQR; moreover, our study highlights that in order to reach a broader understanding of potential heterogeneous associations between height-for-age and its driving factors, one needs to consider the impact not only on lower, but also on higher quantiles.

Table 4: Instrumental variable quantile regressions of the HAZ score

	HAZ Instrumental variable quantile regressions								
	Q = 0.025 (M ₈)	Q = 0.05 (M ₉)	Q = 0.10 (M ₁₀)	Q = 0.25 (M ₁₁)	Q = 0.50 (M ₁₂)	Q = 0.75 (M ₁₃)	Q = 0.90 (M ₁₄)	Q = 0.95 (M ₁₅)	Q = 0.975 (M ₁₆)
Intercept	-9.156*** (0.964)	-10.116*** (0.774)	-10.422*** (0.590)	-9.996*** (0.429)	-9.250*** (0.414)	-8.645*** (0.466)	-7.360*** (0.668)	-5.612*** (1.035)	-4.361*** (1.290)
A. Household characteristics (X_H)									
Household is using primarily solid fuel	-0.135 (0.317)	-0.120 (0.181)	-0.025 (0.136)	-0.105 (0.096)	-0.165* (0.090)	-0.170* (0.095)	-0.320** (0.154)	-0.225 (0.255)	-0.435 (0.277)
Kitchen is separate	0.122 (0.093)	0.066 (0.061)	0.112** (0.049)	0.110*** (0.035)	0.143*** (0.032)	0.095** (0.038)	0.093* (0.055)	0.010 (0.079)	-0.002 (0.115)
House has any windows	0.150* (0.090)	0.130* (0.077)	0.126** (0.053)	0.131*** (0.040)	0.092** (0.039)	0.093** (0.043)	0.103* (0.062)	0.057 (0.088)	-0.035 (0.117)
Rural area	0.323** (0.144)	0.261*** (0.087)	0.096 (0.068)	0.092** (0.046)	0.033 (0.042)	0.062 (0.049)	0.065 (0.076)	0.128 (0.113)	0.233* (0.133)
Nr. household members	-0.024 (0.015)	-0.017* (0.009)	-0.009 (0.007)	-0.011** (0.005)	-0.010** (0.005)	-0.012** (0.005)	-0.010 (0.008)	0.002 (0.014)	0.021 (0.017)
Floor not covered	0.064 (0.097)	0.008 (0.078)	0.035 (0.063)	0.023 (0.047)	0.011 (0.042)	0.028 (0.049)	0.134** (0.067)	0.109 (0.114)	0.241 (0.162)
Open defecation	-0.127 (0.108)	-0.135* (0.073)	-0.095 (0.063)	-0.091** (0.046)	0.010 (0.045)	-0.033 (0.051)	-0.034 (0.075)	-0.056 (0.111)	-0.011 (0.164)
B. Child's characteristics (X_C)									
Child's age (months)	-0.043*** (0.005)	-0.043*** (0.004)	-0.042*** (0.003)	-0.041*** (0.002)	-0.040*** (0.002)	-0.040*** (0.002)	-0.038*** (0.003)	-0.037*** (0.004)	-0.032*** (0.006)
Male child	-0.008 (0.068)	-0.070 (0.050)	-0.056 (0.040)	-0.006 (0.029)	0.035 (0.027)	-0.028 (0.031)	0.033 (0.047)	0.060 (0.065)	-0.055 (0.098)
Birth order	-0.027 (0.021)	-0.039** (0.018)	-0.044*** (0.012)	-0.030*** (0.011)	-0.018* (0.009)	-0.013 (0.011)	0.029* (0.017)	0.059** (0.027)	0.090** (0.038)
Diverse diet	0.425*** (0.084)	0.350*** (0.057)	0.400*** (0.044)	0.448*** (0.031)	0.439*** (0.030)	0.449*** (0.035)	0.499*** (0.058)	0.394*** (0.084)	0.359*** (0.114)
Breastfeeding according to WHO	0.258*** (0.100)	0.110 (0.078)	-0.041 (0.058)	-0.095** (0.040)	-0.065* (0.038)	-0.059 (0.046)	-0.237*** (0.077)	-0.374*** (0.110)	-0.514*** (0.149)
Had fever in the past two weeks	-0.032 (0.093)	-0.048 (0.072)	-0.062 (0.056)	0.021 (0.038)	0.009 (0.036)	-0.064* (0.038)	-0.053 (0.064)	-0.106 (0.090)	-0.029 (0.151)
Had diarrhea in the past two weeks	-0.140 (0.116)	-0.021 (0.076)	-0.004 (0.054)	-0.126*** (0.042)	-0.119*** (0.038)	-0.164*** (0.045)	-0.219*** (0.062)	-0.343*** (0.083)	-0.407*** (0.125)
C. Parents' characteristics (X_P)									
Mother's BMI	0.038*** (0.010)	0.033*** (0.007)	0.034*** (0.006)	0.041*** (0.004)	0.037*** (0.004)	0.042*** (0.005)	0.048*** (0.008)	0.055*** (0.012)	0.053*** (0.019)
Mother's height	0.029*** (0.006)	0.040*** (0.005)	0.046*** (0.004)	0.049*** (0.003)	0.051*** (0.003)	0.054*** (0.003)	0.052*** (0.004)	0.047*** (0.006)	0.042*** (0.008)
Nr. years of mother's education	0.030*** (0.011)	0.032*** (0.008)	0.036*** (0.006)	0.023*** (0.004)	0.013*** (0.004)	0.008* (0.004)	0.001 (0.007)	0.004 (0.009)	0.004 (0.013)
Mother is currently working	0.109 (0.074)	0.055 (0.060)	-0.004 (0.047)	-0.055 (0.035)	-0.065** (0.033)	-0.074** (0.038)	-0.080 (0.062)	-0.081 (0.082)	-0.106 (0.120)
Mother took iron during pregnancy	0.139* (0.078)	0.099 (0.063)	0.105** (0.049)	0.089** (0.037)	0.048 (0.034)	-0.018 (0.040)	-0.087 (0.059)	-0.144 (0.089)	-0.204* (0.115)
Mother is smoking	0.031 (0.107)	0.002 (0.089)	-0.046 (0.065)	-0.052 (0.050)	-0.039 (0.046)	-0.109** (0.050)	-0.126 (0.080)	-0.213* (0.112)	-0.102 (0.214)
Household assets	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
State fixed effects	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Nr. observations	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659

Notes: The table captures estimation results from instrumental variables quantile regressions models of the height-for-age score. Errors were clustered by household, and robust standard errors are presented in parentheses below each coefficient. Significance levels are indicated by ***, **, *, marking the 1%, 5%, and 10% levels, respectively.

7 Discussion and conclusions

The detrimental effects of indoor air pollution on children's health are becoming increasingly apparent, particularly on children living in India, where the prevalence of solid fuel use is extensive.

We analyse the most recent data available from the 2005-2006 India's National Family Health Survey (NFHS-3) in order to determine what are the main drivers of stunting among Indian children. In particular, we test whether the use of solid fuels for cooking and heating is linked to the prevalence of stunting in children, and estimate the extent of the impact after controlling for other confounding factors.

Stunting is not only a proxy for temporary health problems, but it indicates chronic malnutrition and life-long wellbeing deficiencies. It has been shown to be associated with lower educational outcomes and diminished life accomplishments. We believe it is, thus, highly important to understand its drivers and implement policies to largely eliminate it.

Our analysis brings strong evidence that the exposure to solid fuel smoke leads to a lower height-for-age ratio and can explain the prevalence of stunting and severe stunting among Indian children. Other contributing factors are an urban setting, and the practice of openair defecation. Stunting rates appear to be larger for children of older age, of higher birth order, and who have been sick recently. Parental characteristics appear highly relevant for the prevalence of stunting, such as mother's height and body mass index. Moreover, stunted children are associated to mothers with lower levels of education.

Our results reveal that households where indoor air pollution is likely to have lower concentrations, due to the existence of ventilation options (such as windows or a separate room for kitchen) tend to have a lower prevalence of stunting. Diffusing ventilation practices among households of all income levels, either through awareness campaigns or by facilitating access, might be a high potential - low cost measure to reduce the prevalence of stunting among children.

Comparing our work with studies based on data from earlier surveys in India shows that (i) the prevalence of stunting did not decrease during the last decade, and that (ii) one of the main contributing factors remains the exposure to solid fuel smoke. The large share of stunted children remained stable over time despite high average economic growth in India and a gradual expansion of access to grid electricity across the country. One explanation for this unsatisfactory development outcome is that, although incomes and household wealth might have increased over time, families have a preference for which household goods or habits to change first. Changing cooking practices (towards cleaner fuels or more efficient cookstoves) might be perceived as secondary objectives when households climb the income ladder. Increasing awareness with regard to the long-term health and environmental damages of burning solid fuels might help households re-prioritize their acquisitions or at least convince them to make use of ventilation practices to reduce their exposure to indoor air

pollution.

The consequences of burning solid fuels for cooking, heating, and lighting go beyond the health condition of the families that use them. Besides increasing indoor air pollution and particulate matter concentrations, burning solid fuels results in CO₂ emissions and black carbon (BC) particles that have large radiative forcing capacities and contribute to climate change. This pollution impact with global consequences can be seen, however, as a point of opportunity for change. Specifically, researchers and policy makers aware of the double impact are currently attempting to include BC issues into the international negotiations agenda ([Grieshop et al., 2009](#)). Facilitating the adoption of efficient cookstoves, and other home appliances that would require less burning of solid fuel, would achieve the double benefit of avoiding health problems (such as stunting) among households in developing countries and mitigating climate change for both the developed and developing world ([Kar et al., 2012](#)).

A Appendix

A.1 The endogeneity of household fuel choice

Table A.1: Regression models of predicted stunting, severe stunting, and height-for-age score on the fuel choice indicator.

	Stunting $1_{\text{HAZ} < -2}$ (\mathcal{M}_{A1})	Severe stunting $1_{\text{HAZ} < -3}$ (\mathcal{M}_{A2})	HAZ HAZ (\mathcal{M}_{A3})
Intercept	0.242*** (0.002)	0.083*** (0.001)	-0.990*** (0.010)
Household is using primarily solid fuel	0.151*** (0.003)	0.091*** (0.002)	-0.562*** (0.012)
Nr. observations	14,659	14,659	14,659
Adj. R ²	0.140	0.137	0.129

Notes: The table capture estimation results from linear models of stunting (\mathcal{M}_{A1}) and severe stunting (\mathcal{M}_{A2}), OLS model for the HAZ score (\mathcal{M}_{A3}). We follow the procedure described in [Chay and Greenstone \(2003\)](#). In a first step, not presented here, we predict the outcome variable on all exogenous variables in the structural equation (Eq. 6), excluding fuel choice. In a second step, presented here, we regress the predicted outcome variable values on fuel choice. Errors were clustered by household, and robust standard errors are presented in parentheses below each coefficient. Significance levels are indicated by ***, **, *, marking the 1%, 5% and 10% levels, respectively.

A.2 OLS and IV supporting results

Table A.2: Regression models of stunting, severe stunting, HAZ and solid fuel: Estimated coefficients of asset variables for \mathcal{M}_1 - \mathcal{M}_7 in Section 6.

	LPM		OLS	IV			IV first stage
	$1_{\text{HAZ} < -2}$ (\mathcal{M}_{A4})	$1_{\text{HAZ} < -3}$ (\mathcal{M}_{A5})	HAZ (\mathcal{M}_{A6})	$1_{\text{HAZ} < -2}$ (\mathcal{M}_{A7})	$1_{\text{HAZ} < -3}$ (\mathcal{M}_{A8})	HAZ (\mathcal{M}_{A9})	Solid fuel (\mathcal{M}_{A10})
Electricity	-0.012 (0.016)	-0.002 (0.013)	-0.024 (0.053)	-0.012 (0.016)	-0.003 (0.013)	-0.021 (0.053)	0.01 (0.01)
Radio	-0.007 (0.011)	-0.003 (0.009)	0.014 (0.037)	-0.007 (0.011)	-0.003 (0.009)	0.014 (0.037)	0.00 (0.01)
TV	0.014 (0.014)	-0.015 (0.010)	0.047 (0.043)	0.014 (0.014)	-0.013 (0.011)	0.041 (0.043)	-0.02*** (0.01)
Phone	0.009 (0.017)	0.017 (0.011)	-0.038 (0.056)	0.010 (0.017)	0.018* (0.011)	-0.043 (0.056)	-0.03** (0.01)
Computer	-0.029 (0.029)	0.014 (0.018)	0.155 (0.121)	-0.028 (0.029)	0.017 (0.018)	0.146 (0.121)	0.01 (0.02)
Fridge	-0.027 (0.017)	0.011 (0.011)	0.136** (0.056)	-0.025 (0.017)	0.016 (0.011)	0.118** (0.057)	-0.07*** (0.01)
Bicycle	-0.022** (0.011)	-0.026*** (0.009)	0.046 (0.036)	-0.022** (0.011)	-0.026*** (0.009)	0.047 (0.035)	0.01 (0.01)
Motorcycle	-0.004 (0.015)	-0.012 (0.010)	0.044 (0.046)	-0.004 (0.015)	-0.010 (0.010)	0.037 (0.046)	-0.04*** (0.01)
Car	-0.001 (0.024)	0.013 (0.016)	0.266** (0.108)	-0.001 (0.024)	0.014 (0.016)	0.264** (0.108)	0.01 (0.02)
Thresher	-0.074** (0.037)	-0.020 (0.026)	0.119 (0.118)	-0.074** (0.037)	-0.021 (0.026)	0.123 (0.118)	0.03 (0.02)
Tractor	-0.064* (0.035)	-0.023 (0.022)	0.116 (0.113)	-0.066* (0.035)	-0.027 (0.022)	0.130 (0.114)	0.05** (0.02)
Mattress	0.008 (0.013)	0.007 (0.011)	-0.001 (0.042)	0.008 (0.013)	0.008 (0.011)	-0.003 (0.042)	-0.01** (0.01)
Bed	0.033** (0.017)	0.013 (0.013)	-0.126** (0.054)	0.033** (0.017)	0.013 (0.013)	-0.126** (0.054)	-0.00 (0.01)
Chair	-0.005 (0.015)	-0.028** (0.012)	0.114** (0.047)	-0.006 (0.015)	-0.030** (0.012)	0.118** (0.047)	0.01 (0.01)
Table	-0.009 (0.015)	0.005 (0.011)	-0.030 (0.047)	-0.008 (0.015)	0.006 (0.011)	-0.033 (0.047)	-0.02** (0.01)
Pressure cooker	0.003 (0.015)	0.007 (0.012)	-0.020 (0.046)	0.006 (0.016)	0.013 (0.012)	-0.041 (0.049)	-0.10*** (0.01)
Electric fan	-0.035** (0.016)	-0.012 (0.012)	0.082 (0.051)	-0.035** (0.016)	-0.012 (0.012)	0.082 (0.051)	0.01 (0.01)
Sewing machine	0.000 (0.014)	-0.000 (0.010)	0.010 (0.045)	-0.000 (0.014)	-0.001 (0.010)	0.013 (0.045)	-0.00 (0.01)
Water pump	-0.008 (0.016)	-0.007 (0.012)	0.063 (0.054)	-0.007 (0.016)	-0.006 (0.012)	0.060 (0.054)	-0.03*** (0.01)
Bank account	0.012 (0.016)	0.007 (0.011)	-0.106** (0.050)	0.013 (0.016)	0.008 (0.011)	-0.111** (0.049)	-0.02* (0.01)
Nr. observations	14659.000	14659.000	14659.000	14659.000	14659.000	14659.000	
R ²	0.157	0.106	0.207	0.157	0.105	0.207	
Adj. R ²	0.153	0.101	0.203	0.153	0.101	0.203	
First stage <i>F</i> -statistic							1,641.11
Prob > <i>F</i>							0.0000

Clustered, by household, standard errors in parentheses

***, **, *: Significant at the 1%, 5% and 10% levels, respectively.

Table A.3: Logistic regression models of stunting and severe stunting

	Models of stunting		Models of severe stunting	
	$1_{\text{HAZ} < -2}$		$1_{\text{HAZ} < -3}$	
	(\mathcal{M}_{A11})	(\mathcal{M}_{A12})	(\mathcal{M}_{A13})	(\mathcal{M}_{A14})
Intercept	10.220*** (0.612)		7.889*** (0.802)	
A. Household characteristics (X_H)				
Household is using primarily solid fuel	0.215*** (0.066)	0.045*** (0.014)	0.194** (0.093)	0.017** (0.008)
Kitchen is separate	-0.172*** (0.045)	-0.037*** (0.010)	-0.172*** (0.059)	-0.016*** (0.005)
House has any windows	-0.124** (0.050)	-0.027** (0.011)	-0.223*** (0.064)	-0.021*** (0.006)
Rural area	-0.092 (0.056)	-0.020 (0.012)	-0.141* (0.077)	-0.013* (0.007)
Nr. of household members	0.019** (0.007)	0.004** (0.002)	0.018* (0.010)	0.002* (0.001)
Floor not covered	-0.016 (0.058)	-0.003 (0.012)	-0.140* (0.077)	-0.013* (0.007)
Open defecation	0.074 (0.060)	0.016 (0.013)	0.132 (0.081)	0.012 (0.007)
Bank account	0.092 (0.067)	0.020 (0.014)	0.117 (0.096)	0.011 (0.009)
B. Child's characteristics (X_P)				
Child's age (months)	0.044*** (0.002)	0.009*** (0.000)	0.047*** (0.003)	0.004*** (0.000)
Male child	-0.038 (0.038)	-0.008 (0.008)	-0.027 (0.051)	-0.002 (0.005)
Birth order	0.020* (0.012)	0.004* (0.003)	0.030** (0.015)	0.003** (0.001)
Varied food	-0.610*** (0.044)	-0.130*** (0.009)	-0.599*** (0.064)	-0.054*** (0.006)
Breastfeeding according to WHO	0.062 (0.055)	0.013 (0.012)	0.120 (0.074)	0.011 (0.007)
Child had fever in the past two weeks	0.033 (0.051)	0.007 (0.011)	-0.086 (0.069)	-0.008 (0.006)
Child had diarrhea in the past two weeks	0.138** (0.055)	0.029** (0.012)	0.166** (0.073)	0.015** (0.007)
C. Parents' characteristics (X_P)				
Mother's BMI	-0.057*** (0.007)	-0.012*** (0.001)	-0.052*** (0.010)	-0.005*** (0.001)
Mother's height	-0.068*** (0.004)	-0.014*** (0.001)	-0.059*** (0.005)	-0.005*** (0.000)
Nr. years of mother's education	-0.026*** (0.006)	-0.006*** (0.001)	-0.040*** (0.008)	-0.004*** (0.001)
Mother is currently working	0.093** (0.045)	0.020** (0.010)	-0.001 (0.058)	-0.000 (0.005)
Mother took iron during pregnancy	-0.023 (0.045)	-0.005 (0.010)	-0.175*** (0.058)	-0.016*** (0.005)
Mother is smoking	0.003 (0.062)	0.001 (0.013)	0.000 (0.077)	0.000 (0.007)
Household assets	Yes	Yes	Yes	Yes
State fixed effects	Yes	Yes	Yes	Yes
Observations	14,658	14,658	14,658	14,658

Notes: Errors were clustered by household, and robust standard errors are presented in parentheses below each coefficient. Significance levels are indicated by ***, **, *, marking the 1%, 5% and 10% levels, respectively.

Table A.4: Regression of stunting, severe stunting, and HAZ score, with ventilation options interactions.

Notes: The table captures estimation results from linear probability and ordinary least squares models of the three outcome variables. Errors were clustered by household, and robust standard errors are presented in parentheses below each coefficient. Significance levels are indicated by ***, **, *, marking the 1%, 5%, and 10% levels, respectively.

	Interaction (Solid fuel × Separate kitchen)			Interaction (Solid fuel × House has any windows)		
	LPM		OLS	LPM		OLS
	$1_{HAZ < -2}$ (\mathcal{M}_{A15})	$1_{HAZ < -3}$ (\mathcal{M}_{A16})	HAZ (\mathcal{M}_{A17})	$1_{HAZ < -2}$ (\mathcal{M}_{A18})	$1_{HAZ < -3}$ (\mathcal{M}_{A19})	HAZ (\mathcal{M}_{A20})
Intercept	2.279*** (0.147)	1.229*** (0.117)	-8.710*** (0.476)	2.267*** (0.149)	1.236*** (0.119)	-8.609*** (0.482)
A. Household characteristics (X_H)						
Household is using primarily solid fuel	0.041* (0.024)	0.032* (0.018)	-0.099 (0.075)	0.057* (0.033)	0.017 (0.026)	-0.232** (0.112)
Kitchen is separate	-0.046** (0.022)	0.002 (0.015)	0.125* (0.069)	-0.042*** (0.012)	-0.016* (0.009)	0.090** (0.037)
Household is using primarily solid fuel × Kitchen is separate	0.006 (0.025)	-0.022 (0.018)	-0.046 (0.078)			
House has any windows	-0.034** (0.013)	-0.030*** (0.011)	0.135*** (0.042)	-0.021 (0.031)	-0.030 (0.024)	0.030 (0.105)
Household is using primarily solid fuel × House has any windows				-0.014 (0.033)	0.000 (0.026)	0.119 (0.112)
Rural area	-0.021 (0.014)	-0.018* (0.011)	0.075 (0.047)	-0.021 (0.014)	-0.018* (0.011)	0.074 (0.047)
Nr. household members	0.003* (0.002)	0.000 (0.001)	-0.010* (0.006)	0.003* (0.002)	0.000 (0.001)	-0.011* (0.006)
Floor not covered	0.005 (0.015)	0.004 (0.012)	0.018 (0.049)	0.005 (0.015)	0.005 (0.012)	0.023 (0.049)
Open defecation	0.017 (0.016)	0.009 (0.012)	-0.028 (0.049)	0.016 (0.016)	0.010 (0.012)	-0.026 (0.049)
Electricity	-0.012 (0.016)	-0.002 (0.013)	-0.023 (0.053)	-0.012 (0.016)	-0.002 (0.013)	-0.026 (0.053)
Radio	-0.007 (0.011)	-0.003 (0.009)	0.014 (0.037)	-0.007 (0.011)	-0.003 (0.009)	0.013 (0.037)
TV	0.014 (0.014)	-0.015 (0.010)	0.047 (0.043)	0.014 (0.014)	-0.015 (0.010)	0.047 (0.043)
Phone	0.010 (0.017)	0.016 (0.011)	-0.040 (0.056)	0.009 (0.017)	0.017 (0.011)	-0.035 (0.056)
Computer	-0.029 (0.029)	0.013 (0.017)	0.153 (0.121)	-0.029 (0.029)	0.014 (0.018)	0.157 (0.122)
Fridge	-0.027 (0.017)	0.009 (0.011)	0.131** (0.057)	-0.028* (0.017)	0.011 (0.011)	0.140** (0.056)
Bicycle	-0.022** (0.011)	-0.026*** (0.009)	0.046 (0.036)	-0.022** (0.011)	-0.026*** (0.009)	0.046 (0.036)
Motorcycle	-0.004 (0.015)	-0.013 (0.010)	0.042 (0.046)	-0.005 (0.015)	-0.012 (0.010)	0.045 (0.046)
Car	-0.001 (0.024)	0.012 (0.016)	0.264** (0.108)	-0.001 (0.024)	0.013 (0.016)	0.267** (0.108)
Thresher	-0.074** (0.037)	-0.020 (0.026)	0.119 (0.118)	-0.073** (0.037)	-0.020 (0.026)	0.117 (0.118)
Tractor	-0.064* (0.035)	-0.022 (0.022)	0.118 (0.113)	-0.064* (0.035)	-0.023 (0.022)	0.114 (0.113)
Mattress	0.008 (0.013)	0.007 (0.011)	-0.001 (0.042)	0.008 (0.013)	0.007 (0.011)	-0.001 (0.042)
Bed	0.033** (0.017)	0.012 (0.013)	-0.127** (0.055)	0.032* (0.017)	0.013 (0.013)	-0.124** (0.054)
Chair	-0.005 (0.015)	-0.028** (0.012)	0.114** (0.047)	-0.005 (0.015)	-0.028** (0.012)	0.113** (0.047)
Table	-0.009 (0.015)	0.005 (0.011)	-0.031 (0.047)	-0.009 (0.015)	0.005 (0.011)	-0.031 (0.047)
Pressure cooker	0.003 (0.015)	0.007 (0.012)	-0.019 (0.046)	0.003 (0.015)	0.007 (0.012)	-0.021 (0.046)
Electric fan	-0.035** (0.016)	-0.012 (0.012)	0.083 (0.051)	-0.035** (0.016)	-0.012 (0.012)	0.082 (0.051)
Sewing machine	0.000 (0.014)	-0.001 (0.010)	0.009 (0.045)	0.000 (0.014)	-0.000 (0.010)	0.010 (0.045)
Water pump	-0.008 (0.016)	-0.007 (0.012)	0.063 (0.054)	-0.008 (0.016)	-0.007 (0.012)	0.063 (0.054)
Bank account	0.012 (0.016)	0.006 (0.011)	-0.106** (0.050)	0.012 (0.016)	0.007 (0.011)	-0.106** (0.050)
B. Child's characteristics (X_C)						
Child's age	0.009*** (0.001)	0.006*** (0.000)	-0.042*** (0.002)	0.009*** (0.001)	0.006*** (0.000)	-0.042*** (0.002)
Male child	-0.020** (0.010)	-0.013* (0.008)	0.050 (0.031)	-0.019** (0.010)	-0.013* (0.008)	0.050 (0.031)
Birth order	0.002 (0.003)	0.006** (0.003)	-0.001 (0.010)	0.002 (0.003)	0.006** (0.003)	-0.001 (0.010)
Diverse diet	-0.137*** (0.016)	-0.062*** (0.011)	0.467*** (0.050)	-0.137*** (0.016)	-0.062*** (0.011)	0.466*** (0.050)

Table A.4 – continued from previous page

	Interaction (Solid fuel × Separate kitchen)			Interaction (Solid fuel × House has any windows)		
	LPM		OLS	LPM		OLS
	$1_{\text{HAZ} < -2}$ (\mathcal{M}_{A15})	$1_{\text{HAZ} < -3}$ (\mathcal{M}_{A16})	HAZ (\mathcal{M}_{A17})	$1_{\text{HAZ} < -2}$ (\mathcal{M}_{A18})	$1_{\text{HAZ} < -3}$ (\mathcal{M}_{A19})	HAZ (\mathcal{M}_{A20})
Breastfeeding according to WHO	(0.011) 0.028*	(0.008) 0.039***	(0.036) -0.197***	(0.011) 0.028*	(0.008) 0.039***	(0.036) -0.197***
Had fever in the past two weeks	(0.015) 0.008	(0.012) -0.004	(0.051) -0.065	(0.015) 0.008	(0.012) -0.004	(0.051) -0.065
Had diarrhea in the past two weeks	(0.013) 0.032**	(0.011) 0.013	(0.041) -0.152***	(0.013) 0.032**	(0.011) 0.013	(0.041) -0.153***
	(0.014)	(0.011)	(0.043)	(0.014)	(0.011)	(0.043)
C. Parents' characteristics (X_P)						
Mother's BMI	-0.009*** (0.002)	-0.004*** (0.001)	0.039*** (0.005)	-0.010*** (0.002)	-0.004*** (0.001)	0.039*** (0.005)
Mother's height	-0.012*** (0.001)	-0.007*** (0.001)	0.047*** (0.003)	-0.012*** (0.001)	-0.007*** (0.001)	0.047*** (0.003)
Nr. years of mother's education	-0.005*** (0.001)	-0.004*** (0.001)	0.016*** (0.005)	-0.006*** (0.001)	-0.004*** (0.001)	0.017*** (0.005)
Mother is currently working	0.014 (0.012)	0.005 (0.010)	-0.041 (0.039)	0.014 (0.012)	0.005 (0.010)	-0.040 (0.039)
Mother took iron during pregnancy	-0.007 (0.012)	-0.031*** (0.010)	0.050 (0.039)	-0.007 (0.012)	-0.031*** (0.010)	0.049 (0.039)
Mother is smoking	0.009 (0.018)	-0.008 (0.015)	-0.064 (0.056)	0.009 (0.018)	-0.008 (0.015)	-0.063 (0.056)
State fixed effects	Yes	Yes	Yes	Yes	Yes	Yes
Nr. observations	14,659	14,659	14,659	14,659	14,659	14,659
R ²	0.157	0.106	0.207	0.157	0.106	0.207
Adj. R ²	0.153	0.101	0.203	0.153	0.101	0.203

Table A.5: Regression of stunting, severe stunting, and HAZ score, with ventilation options interactions.

Notes: The table captures estimation results from linear probability and ordinary least squares models of the three outcome variables. Errors were clustered by household, and robust standard errors are presented in parentheses below each coefficient. Significance levels are indicated by ***, **, *, and 10% levels, respectively.

	X = 6 months			X = 12 months			X = 18 months			X = 24 months			X = 30 months		
	LPM		OLS	LPM		OLS	LPM		OLS	LPM		OLS	LPM		OLS
	$1_{HAZ < -2}$ ($M_{A,20}$)	$1_{HAZ < -3}$ ($M_{A,21}$)	HAZ ($M_{A,22}$)	$1_{HAZ < -2}$ ($M_{A,23}$)	$1_{HAZ < -3}$ ($M_{A,24}$)	HAZ ($M_{A,25}$)	$1_{HAZ < -2}$ ($M_{A,26}$)	$1_{HAZ < -3}$ ($M_{A,27}$)	HAZ ($M_{A,28}$)	$1_{HAZ < -2}$ ($M_{A,29}$)	$1_{HAZ < -3}$ ($M_{A,30}$)	HAZ ($M_{A,31}$)	$1_{HAZ < -2}$ ($M_{A,32}$)	$1_{HAZ < -3}$ ($M_{A,33}$)	HAZ ($M_{A,34}$)
Intercept	2.328*** (0.147)	1.311*** (0.117)	-8.533*** (0.484)	2.355*** (0.145)	1.317*** (0.116)	-8.886*** (0.471)	2.434*** (0.147)	1.345*** (0.117)	-9.326*** (0.479)	2.460*** (0.148)	1.350*** (0.118)	-9.471*** (0.487)	2.452*** (0.148)	1.349*** (0.118)	-9.460*** (0.487)
Household is using primarily solid fuel	-0.093*** (0.024)	-0.071*** (0.015)	0.042 (0.101)	-0.048** (0.020)	-0.053*** (0.014)	-0.002 (0.073)	-0.015 (0.019)	-0.029** (0.014)	0.023 (0.064)	0.020 (0.018)	0.004 (0.014)	-0.043 (0.060)	0.030* (0.018)	0.008 (0.013)	-0.079 (0.056)
Child's age $\geq X$	0.111*** (0.021)	0.034*** (0.013)	-0.909*** (0.089)	0.148*** (0.017)	0.058*** (0.012)	-0.847*** (0.064)	0.060*** (0.017)	0.031*** (0.012)	-0.401*** (0.058)	-0.030 (0.019)	0.002 (0.013)	0.010 (0.062)	0.002 (0.024)	0.027 (0.018)	-0.147* (0.085)
(Solid fuel) \times (Child's age $\geq X$)	0.157*** (0.023)	0.101*** (0.014)	-0.187* (0.098)	0.141*** (0.020)	0.105*** (0.014)	-0.200** (0.072)	0.126*** (0.021)	0.098*** (0.015)	-0.322*** (0.068)	0.077*** (0.023)	0.041** (0.017)	-0.271*** (0.076)	0.091*** (0.030)	0.055** (0.024)	-0.299*** (0.103)
A. Household characteristics (X_{it})															
Kitchen is separate	-0.042*** (0.012)	-0.016 (0.010)	0.090** (0.038)	-0.043*** (0.012)	-0.016* (0.009)	0.093** (0.037)	-0.042*** (0.012)	-0.016* (0.009)	0.089** (0.038)	-0.040*** (0.012)	-0.015 (0.010)	0.082** (0.039)	-0.040*** (0.012)	-0.015 (0.010)	0.080** (0.039)
House has any windows	-0.030** (0.013)	-0.028** (0.011)	0.116*** (0.043)	-0.036*** (0.013)	-0.031*** (0.011)	0.137*** (0.042)	-0.035*** (0.013)	-0.031*** (0.011)	0.138*** (0.042)	-0.031** (0.013)	-0.029*** (0.011)	0.125*** (0.043)	-0.031** (0.013)	-0.028*** (0.011)	0.124*** (0.043)
Rural area	-0.018 (0.014)	-0.017 (0.011)	0.063 (0.048)	-0.019 (0.014)	-0.018 (0.011)	0.067 (0.047)	-0.021 (0.014)	-0.018* (0.011)	0.075 (0.048)	-0.022 (0.014)	-0.019* (0.011)	0.080** (0.048)	-0.022 (0.014)	-0.019* (0.011)	0.079 (0.048)
Nr. household members	0.002 (0.002)	-0.000 (0.001)	-0.007 (0.006)	0.003* (0.002)	0.000 (0.001)	-0.009* (0.005)	0.002 (0.002)	0.000 (0.001)	0.000 (0.006)	0.002 (0.002)	-0.000 (0.001)	-0.005 (0.006)	0.002 (0.002)	-0.001 (0.001)	-0.005 (0.006)
Floor not covered	0.008 (0.016)	0.006 (0.012)	0.005 (0.049)	0.007 (0.015)	0.005 (0.012)	0.011 (0.048)	-0.000 (0.015)	-0.000 (0.012)	0.040 (0.049)	0.003 (0.016)	0.003 (0.012)	0.031 (0.050)	0.003 (0.016)	0.004 (0.012)	0.030 (0.050)
Open defecation	0.017 (0.016)	0.009 (0.012)	-0.029 (0.050)	0.021 (0.015)	0.012 (0.012)	-0.040 (0.048)	0.013 (0.016)	0.008 (0.012)	-0.012 (0.049)	0.014 (0.016)	0.008 (0.012)	-0.017 (0.050)	0.013 (0.016)	0.007 (0.012)	-0.009 (0.050)
Electricity	-0.013 (0.016)	-0.003 (0.013)	-0.019 (0.054)	-0.012 (0.016)	-0.002 (0.013)	-0.021 (0.052)	-0.010 (0.016)	-0.000 (0.013)	-0.031 (0.054)	-0.013 (0.017)	-0.003 (0.014)	-0.019 (0.055)	-0.014 (0.017)	-0.003 (0.013)	-0.014 (0.054)
Radio	-0.011 (0.011)	-0.005 (0.009)	0.028 (0.037)	-0.005 (0.011)	-0.002 (0.009)	0.005 (0.036)	-0.007 (0.011)	-0.003 (0.009)	0.012 (0.037)	-0.007 (0.011)	-0.003 (0.009)	0.011 (0.038)	-0.007 (0.011)	-0.003 (0.009)	0.012 (0.038)
TV	0.015 (0.014)	-0.014 (0.011)	0.044 (0.044)	0.012 (0.014)	-0.016 (0.011)	0.051 (0.044)	0.015 (0.014)	-0.014 (0.011)	0.037 (0.044)	0.014 (0.014)	-0.015 (0.011)	0.044 (0.044)	0.014 (0.014)	-0.015 (0.011)	0.044 (0.044)
Phone	0.006 (0.016)	0.015 (0.011)	-0.013 (0.056)	0.011 (0.016)	0.018 (0.011)	-0.040 (0.055)	0.010 (0.016)	0.017 (0.011)	-0.039 (0.056)	0.008 (0.017)	0.016 (0.011)	-0.035 (0.057)	0.009 (0.017)	0.016 (0.011)	-0.035 (0.057)
Computer	-0.034 (0.028)	0.010 (0.017)	0.174 (0.122)	-0.029 (0.028)	0.013 (0.017)	0.149 (0.118)	-0.027 (0.028)	0.015 (0.017)	0.152 (0.122)	-0.033 (0.028)	0.012 (0.017)	0.171 (0.124)	-0.033 (0.028)	0.011 (0.017)	0.176 (0.124)
Fridge	-0.027 (0.017)	0.011 (0.011)	0.136** (0.056)	-0.027 (0.017)	0.011 (0.011)	0.133** (0.055)	-0.026 (0.017)	0.012 (0.011)	0.134** (0.056)	-0.029* (0.017)	0.010 (0.011)	0.142** (0.058)	-0.029* (0.017)	0.010 (0.011)	0.141** (0.058)
Bicycle	-0.020* (0.011)	-0.025*** (0.009)	0.038 (0.036)	-0.023** (0.011)	-0.026*** (0.009)	0.051 (0.035)	-0.021* (0.011)	-0.025*** (0.009)	0.043 (0.036)	-0.021* (0.011)	-0.026*** (0.009)	0.043 (0.036)	-0.020** (0.011)	-0.025*** (0.009)	0.038 (0.036)
Motorcycle	-0.008 (0.015)	-0.015 (0.010)	0.056 (0.046)	-0.006 (0.014)	-0.013 (0.010)	0.051 (0.046)	-0.007 (0.015)	-0.014 (0.010)	0.051 (0.046)	-0.007 (0.015)	-0.014 (0.010)	0.053 (0.046)	-0.007 (0.015)	-0.014 (0.010)	0.054 (0.047)
Car	-0.006 (0.024)	0.009 (0.016)	0.275*** (0.112)	-0.006 (0.024)	0.009 (0.016)	0.275*** (0.104)	-0.009 (0.024)	0.008 (0.016)	0.301*** (0.108)	-0.006 (0.024)	0.010 (0.016)	0.285** (0.111)	-0.005 (0.024)	0.011 (0.016)	0.279** (0.111)
Threshold	-0.073** (0.036)	-0.018 (0.026)	0.119 (0.117)	-0.073** (0.037)	-0.019 (0.026)	0.113 (0.118)	-0.073** (0.037)	-0.020 (0.026)	0.113 (0.121)	-0.067* (0.037)	-0.016 (0.026)	0.089 (0.121)	-0.065* (0.037)	-0.015 (0.026)	0.081 (0.120)
Tractor	-0.067* (0.035)	-0.025 (0.022)	0.129 (0.113)	-0.060* (0.035)	-0.021 (0.022)	0.104 (0.112)	-0.063* (0.036)	-0.022 (0.022)	0.114 (0.117)	-0.064* (0.035)	-0.023 (0.022)	0.118 (0.116)	-0.065* (0.035)	-0.023 (0.022)	0.118 (0.115)

Table A.5 – continued from previous page

	X = 6 months			X = 12 months			X = 18 months			X = 24 months			X = 30 months		
	LPM		OLS	LPM		OLS	LPM		OLS	LPM		OLS	LPM		OLS
	$1_{HAZ < -2}$ ($M_{A,20}$)	$1_{HAZ < -3}$ ($M_{A,21}$)	HAZ ($M_{A,22}$)	$1_{HAZ < -2}$ ($M_{A,23}$)	$1_{HAZ < -3}$ ($M_{A,24}$)	HAZ ($M_{A,25}$)	$1_{HAZ < -2}$ ($M_{A,26}$)	$1_{HAZ < -3}$ ($M_{A,27}$)	HAZ ($M_{A,28}$)	$1_{HAZ < -2}$ ($M_{A,29}$)	$1_{HAZ < -3}$ ($M_{A,30}$)	HAZ ($M_{A,31}$)	$1_{HAZ < -2}$ ($M_{A,32}$)	$1_{HAZ < -3}$ ($M_{A,33}$)	HAZ ($M_{A,34}$)
Mattress	0.003 (0.013)	0.004 (0.011)	0.017 (0.042)	0.003 (0.013)	0.004 (0.011)	0.015 (0.041)	0.005 (0.013)	0.005 (0.011)	0.009 (0.042)	0.005 (0.014)	0.006 (0.011)	0.007 (0.043)	0.006 (0.013)	0.006 (0.011)	0.007 (0.043)
Bed	0.028* (0.017)	0.011 (0.014)	-0.107* (0.055)	0.030* (0.016)	0.011 (0.013)	-0.116** (0.054)	0.033** (0.017)	0.014 (0.013)	-0.134** (0.054)	0.033** (0.017)	0.013 (0.014)	-0.129** (0.055)	0.034** (0.017)	0.014 (0.014)	-0.132** (0.055)
Chair	-0.005 (0.015)	-0.028** (0.012)	0.113** (0.048)	0.002 (0.015)	-0.024** (0.012)	0.086* (0.047)	-0.004 (0.015)	-0.027** (0.012)	0.107** (0.048)	-0.003 (0.015)	-0.028** (0.012)	0.107** (0.048)	-0.004 (0.015)	-0.028** (0.012)	0.111** (0.048)
Table	-0.010 (0.015)	0.004 (0.011)	-0.026 (0.048)	-0.013 (0.015)	0.002 (0.011)	-0.014 (0.047)	-0.012 (0.015)	0.003 (0.011)	-0.018 (0.048)	-0.008 (0.015)	0.006 (0.011)	-0.035 (0.048)	-0.007 (0.015)	-0.041 (0.011)	-0.041 (0.048)
Pressure cooker	0.007 (0.015)	0.010 (0.012)	-0.042 (0.047)	0.006 (0.015)	0.010 (0.012)	-0.037 (0.046)	0.004 (0.015)	0.007 (0.012)	-0.026 (0.047)	0.009 (0.016)	0.011 (0.012)	-0.047 (0.048)	0.008 (0.016)	0.010 (0.012)	-0.043 (0.048)
Electric fan	-0.033** (0.016)	-0.012 (0.013)	0.076 (0.052)	-0.030* (0.016)	-0.010 (0.012)	0.063 (0.051)	-0.037** (0.016)	-0.014 (0.012)	0.091* (0.052)	-0.035** (0.016)	-0.013 (0.013)	0.086* (0.052)	-0.036** (0.016)	-0.013 (0.013)	0.088* (0.052)
Sewing machine	0.003 (0.014)	0.001 (0.010)	-0.004 (0.046)	0.005 (0.014)	0.003 (0.010)	-0.005 (0.045)	0.001 (0.014)	0.000 (0.010)	0.010 (0.045)	-0.001 (0.014)	-0.001 (0.010)	0.018 (0.046)	-0.002 (0.014)	-0.002 (0.010)	0.021 (0.046)
Water pump	-0.010 (0.016)	-0.008 (0.012)	0.071 (0.053)	-0.010 (0.016)	-0.008 (0.012)	0.065 (0.053)	-0.006 (0.016)	-0.005 (0.012)	0.051 (0.055)	-0.007 (0.016)	-0.006 (0.012)	0.060 (0.055)	-0.007 (0.016)	-0.006 (0.012)	0.060 (0.055)
Bank account	0.010 (0.016)	0.007 (0.011)	-0.093* (0.050)	0.011 (0.015)	0.007 (0.011)	-0.102** (0.049)	0.017 (0.016)	0.010 (0.011)	-0.126** (0.050)	0.019 (0.016)	0.010 (0.011)	-0.132** (0.051)	0.018 (0.016)	0.010 (0.011)	-0.131** (0.051)
B. Child's characteristics (X_C)															
Male child	-0.017* (0.010)	-0.011 (0.008)	0.040 (0.032)	-0.019** (0.010)	-0.012 (0.008)	0.050 (0.031)	0.018* (0.010)	-0.012 (0.008)	0.041 (0.032)	0.015 (0.010)	-0.015 (0.008)	0.032 (0.032)	-0.015 (0.010)	-0.011 (0.008)	0.031 (0.032)
Birth order	0.004 (0.003)	0.008*** (0.003)	-0.012 (0.010)	0.003 (0.003)	0.007*** (0.003)	-0.007 (0.010)	0.003 (0.003)	0.007*** (0.003)	-0.006 (0.010)	0.004 (0.003)	0.007*** (0.003)	-0.009 (0.011)	0.004 (0.003)	0.007*** (0.003)	-0.008 (0.011)
Diverse diet	-0.117*** (0.012)	-0.065*** (0.009)	0.377*** (0.040)	-0.098*** (0.011)	-0.049*** (0.008)	0.363*** (0.037)	-0.162*** (0.011)	-0.078*** (0.008)	0.601*** (0.025)**	-0.206*** (0.010)	-0.106*** (0.008)	0.765*** (0.035)	-0.206*** (0.010)	-0.107*** (0.007)	0.772*** (0.035)
Breastfeeding according to WHO	-0.023* (0.014)	0.001 (0.011)	0.016 (0.049)	0.020 (0.015)	0.024** (0.011)	-0.147*** (0.050)	0.008 (0.015)	0.025** (0.012)	-0.102** (0.051)	-0.037** (0.015)	-0.037** (0.012)	0.069 (0.051)	-0.032** (0.014)	0.003 (0.011)	0.060 (0.049)
Had fever in the past two weeks	-0.002 (0.013)	-0.009 (0.011)	-0.020 (0.042)	0.006 (0.013)	-0.005 (0.011)	-0.057 (0.041)	0.006 (0.013)	-0.001 (0.011)	-0.083** (0.042)	0.010 (0.013)	-0.003 (0.011)	-0.072* (0.042)	0.010 (0.013)	-0.003 (0.011)	-0.073* (0.042)
Had diarrhea in the past two weeks	0.012 (0.014)	0.000 (0.011)	-0.061 (0.045)	0.032** (0.014)	0.011 (0.011)	-0.144*** (0.042)	0.030** (0.014)	0.012 (0.011)	-0.140*** (0.044)	0.016 (0.014)	0.004 (0.011)	-0.085* (0.045)	0.017 (0.014)	0.004 (0.011)	-0.088* (0.045)
C. Parents' characteristics (X_P)															
Mother's BMI	-0.008*** (0.002)	-0.003*** (0.001)	0.032*** (0.005)	-0.009*** (0.002)	-0.004*** (0.001)	0.036*** (0.005)	-0.010*** (0.002)	-0.004*** (0.001)	0.039*** (0.005)	-0.009*** (0.002)	-0.004*** (0.001)	0.038*** (0.006)	-0.009*** (0.002)	-0.004*** (0.001)	0.038*** (0.006)
Mother's height	-0.012*** (0.001)	-0.007*** (0.001)	0.046*** (0.003)	-0.013*** (0.001)	-0.007*** (0.001)	0.048*** (0.003)	-0.012*** (0.001)	-0.007*** (0.001)	0.047*** (0.003)	-0.012*** (0.001)	-0.007*** (0.001)	0.045*** (0.003)	-0.012*** (0.001)	-0.007*** (0.001)	0.045*** (0.003)
Nr. years of mother's education	-0.006*** (0.001)	-0.005*** (0.001)	0.019*** (0.005)	-0.006*** (0.001)	-0.004*** (0.001)	0.017*** (0.005)	-0.006*** (0.001)	-0.004*** (0.001)	0.017*** (0.005)	-0.006*** (0.001)	-0.005*** (0.001)	0.018*** (0.005)	-0.006*** (0.001)	-0.005*** (0.001)	0.018*** (0.005)
Mother is currently working	0.025** (0.012)	0.013 (0.010)	-0.089** (0.039)	0.008 (0.012)	0.004 (0.010)	-0.034 (0.038)	0.018 (0.012)	0.006 (0.010)	-0.064* (0.039)	0.033*** (0.012)	0.016 (0.010)	-0.120*** (0.039)	0.032*** (0.012)	0.015 (0.010)	-0.118*** (0.039)
Mother took iron during pregnancy	-0.007 (0.012)	-0.031*** (0.010)	0.050 (0.039)	-0.006 (0.012)	-0.031*** (0.010)	0.045 (0.042)	-0.006 (0.012)	-0.031*** (0.012)	0.048 (0.039)	-0.008 (0.012)	-0.031*** (0.012)	0.055 (0.040)	-0.008 (0.012)	-0.032*** (0.010)	0.056 (0.040)
Mother is smoking	0.007 (0.018)	-0.008 (0.015)	-0.058 (0.057)	0.007 (0.018)	-0.009 (0.015)	-0.058 (0.055)	0.010 (0.018)	-0.008 (0.015)	-0.068 (0.056)	0.012 (0.018)	-0.006 (0.015)	-0.077 (0.057)	0.012 (0.018)	-0.006 (0.015)	-0.078 (0.057)
State fixed effects	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Nr. observations	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659
R ²	0.152	0.095	0.196	0.179	0.112	0.224	0.155	0.105	0.195	0.133	0.088	0.165	0.134	0.090	0.168
Adj. R ²	0.148	0.091	0.193	0.175	0.108	0.220	0.151	0.101	0.191	0.129	0.084	0.161	0.130	0.086	0.164

Table A.6: Quantile regression models of HAZ

Notes: The table captures estimation results from instrumental variables quantile regressions models of the height-for-age score. Errors were clustered by household, and robust standard errors are presented in parentheses below each coefficient. Significance levels are indicated by ***, **, *, marking the 1%, 5%, and 10% levels, respectively.

	HAZ quantile regressions								
	Q = 0.025 (\mathcal{M}_{A35})	Q = 0.05 (\mathcal{M}_{A36})	Q = 0.10 (\mathcal{M}_{A37})	Q = 0.25 (\mathcal{M}_{A38})	Q = 0.50 (\mathcal{M}_{A39})	Q = 0.75 (\mathcal{M}_{A40})	Q = 0.90 (\mathcal{M}_{A41})	Q = 0.95 (\mathcal{M}_{A42})	Q = 0.975 (\mathcal{M}_{A43})
Intercept	-9.299*** (0.544)	-10.251*** (0.490)	-10.535*** (0.511)	-10.149*** (0.360)	-9.479*** (0.409)	-8.693*** (0.429)	-7.660*** (0.503)	-5.973*** (0.654)	-3.778*** (0.827)
A. Household characteristics (X_H)									
Household is using primarily solid fuel	-0.126* (0.075)	-0.030 (0.059)	-0.098* (0.053)	-0.089** (0.040)	-0.089** (0.043)	-0.041 (0.045)	-0.034 (0.050)	-0.080 (0.060)	-0.222*** (0.084)
Kitchen is separate	0.125*** (0.044)	0.071* (0.040)	0.121*** (0.040)	0.112*** (0.030)	0.141*** (0.032)	0.099*** (0.034)	0.088** (0.044)	0.012 (0.057)	-0.012 (0.067)
House has any windows	0.149*** (0.052)	0.124*** (0.043)	0.137*** (0.045)	0.130*** (0.034)	0.087** (0.037)	0.097*** (0.037)	0.120*** (0.043)	0.034 (0.063)	-0.043 (0.075)
Rural area	0.315*** (0.065)	0.229*** (0.054)	0.110** (0.048)	0.089** (0.035)	0.006 (0.038)	0.016 (0.039)	-0.001 (0.045)	0.090 (0.058)	0.166** (0.083)
Nr. household members	-0.024** (0.010)	-0.017*** (0.005)	-0.009 (0.007)	-0.011*** (0.004)	-0.012** (0.005)	-0.013*** (0.004)	-0.012** (0.006)	-0.000 (0.009)	0.024** (0.011)
Floor not covered	0.066 (0.069)	0.001 (0.044)	0.040 (0.048)	0.022 (0.036)	0.008 (0.041)	0.009 (0.043)	0.155*** (0.046)	0.099 (0.067)	0.226*** (0.080)
Open defecation	-0.127** (0.062)	-0.124*** (0.046)	-0.077 (0.051)	-0.092** (0.037)	-0.017 (0.043)	-0.054 (0.045)	-0.057 (0.048)	-0.067 (0.067)	-0.063 (0.090)
Electricity	-0.136** (0.058)	-0.092* (0.052)	0.008 (0.057)	-0.009 (0.044)	0.028 (0.046)	0.038 (0.049)	-0.082 (0.061)	-0.079 (0.085)	-0.249*** (0.090)
Radio	0.055 (0.047)	0.003 (0.035)	0.004 (0.039)	0.027 (0.027)	0.058* (0.030)	-0.005 (0.030)	-0.043 (0.035)	-0.032 (0.055)	-0.203*** (0.064)
TV	0.130** (0.058)	0.115*** (0.044)	0.107** (0.048)	0.048 (0.034)	0.040 (0.038)	0.041 (0.047)	0.053 (0.047)	-0.061 (0.057)	-0.080 (0.068)
Phone	-0.292*** (0.088)	-0.118* (0.067)	-0.096 (0.067)	-0.051 (0.040)	-0.022 (0.047)	0.028 (0.050)	0.074 (0.052)	0.049 (0.059)	-0.129 (0.102)
Computer	0.017 (0.169)	0.037 (0.203)	0.016 (0.091)	0.128** (0.055)	0.142* (0.081)	0.191** (0.081)	0.126 (0.103)	0.028 (0.199)	0.050 (0.363)
Fridge	0.246*** (0.083)	0.164*** (0.054)	0.073 (0.062)	0.074* (0.044)	0.114** (0.046)	0.113** (0.046)	0.178*** (0.054)	0.231*** (0.076)	0.520*** (0.119)
Bicycle	0.128*** (0.048)	0.155*** (0.037)	0.144*** (0.038)	0.091*** (0.027)	0.080*** (0.030)	-0.005 (0.031)	-0.008 (0.039)	-0.039 (0.050)	0.097 (0.071)
Motorcycle	0.001 (0.074)	0.089** (0.044)	0.004 (0.053)	-0.035 (0.035)	0.011 (0.040)	0.069* (0.040)	0.087** (0.043)	0.114* (0.060)	0.209* (0.122)
Car	-0.010 (0.182)	-0.107 (0.148)	0.159* (0.090)	0.120** (0.049)	0.034 (0.074)	0.096 (0.081)	0.099 (0.114)	0.260 (0.221)	0.744 (0.473)
Thresher	0.022 (0.115)	-0.003 (0.331)	0.135 (0.085)	0.147* (0.085)	0.067 (0.099)	0.004 (0.099)	0.098 (0.251)	0.451 (0.849)	0.656*** (0.210)
Tractor	0.174 (0.132)	0.062 (0.086)	0.172 (0.140)	0.096 (0.061)	0.040 (0.160)	0.046 (0.110)	-0.110 (0.113)	-0.153 (0.163)	0.214 (0.377)
Mattress	0.050 (0.052)	0.063 (0.044)	0.038 (0.046)	-0.007 (0.034)	-0.022 (0.038)	0.015 (0.038)	0.063 (0.049)	-0.035 (0.063)	-0.032 (0.076)
Bed	0.098* (0.057)	-0.060 (0.068)	-0.063 (0.052)	-0.139*** (0.045)	-0.062 (0.048)	-0.141*** (0.050)	-0.303*** (0.068)	-0.473*** (0.131)	-0.651*** (0.089)
Chair	0.034 (0.057)	0.114** (0.047)	0.020 (0.046)	0.051 (0.037)	0.018 (0.040)	0.026 (0.041)	0.124*** (0.046)	0.127* (0.068)	0.059 (0.074)
Table	-0.117* (0.063)	-0.053 (0.049)	-0.002 (0.049)	0.067* (0.035)	0.064 (0.040)	0.099** (0.041)	0.021 (0.049)	-0.016 (0.063)	0.094 (0.074)
Pressure cooker	0.195*** (0.056)	0.119** (0.048)	0.060 (0.051)	0.023 (0.036)	0.005 (0.042)	-0.036 (0.044)	-0.024 (0.047)	-0.048 (0.064)	-0.043 (0.080)
Electric fan	0.221*** (0.064)	0.186*** (0.052)	0.102* (0.056)	0.152*** (0.037)	0.079* (0.043)	0.007 (0.045)	0.029 (0.047)	0.209*** (0.068)	0.290*** (0.079)
Sewing machine	0.081 (0.070)	0.087* (0.045)	-0.038 (0.050)	0.010 (0.033)	0.024 (0.037)	0.022 (0.035)	-0.015 (0.044)	-0.015 (0.053)	-0.104 (0.074)
Water pump	0.305*** (0.078)	0.259*** (0.048)	0.148*** (0.049)	0.046 (0.038)	0.027 (0.046)	0.006 (0.048)	0.087 (0.060)	0.036 (0.083)	0.026 (0.103)
Bank account	-0.202*** (0.072)	-0.258*** (0.079)	-0.111* (0.061)	-0.046 (0.038)	-0.034 (0.044)	-0.036 (0.041)	-0.089* (0.047)	-0.048 (0.061)	-0.074 (0.085)
B. Child's characteristics (X_C)									
Child's age (months)	-0.042*** (0.002)	-0.043*** (0.002)	-0.042*** (0.002)	-0.041*** (0.001)	-0.040*** (0.002)	-0.040*** (0.002)	-0.039*** (0.002)	-0.036*** (0.003)	-0.033*** (0.003)
Male child	-0.006 (0.038)	-0.058* (0.033)	-0.064* (0.033)	-0.005 (0.024)	0.033 (0.027)	0.052 (0.027)	0.041 (0.033)	0.041 (0.044)	-0.018 (0.053)
Birth order	-0.026** (0.011)	-0.041*** (0.011)	-0.042*** (0.012)	-0.030*** (0.009)	-0.019** (0.009)	-0.014 (0.010)	0.029** (0.013)	0.061*** (0.019)	0.098*** (0.022)
Diverse diet	0.424*** (0.044)	0.355*** (0.038)	0.410*** (0.038)	0.446*** (0.027)	0.437*** (0.030)	0.446*** (0.031)	0.484*** (0.037)	0.394*** (0.047)	0.371*** (0.060)
Breastfeeding according to WHO	0.259*** (0.057)	0.106 (0.066)	-0.034 (0.054)	-0.096*** (0.039)	-0.055 (0.041)	-0.066 (0.060)	-0.267*** (0.041)	-0.343*** (0.070)	-0.518*** (0.067)
Had fever in the past two weeks	-0.027 (0.045)	-0.070 (0.043)	-0.066 (0.048)	0.016 (0.035)	0.014 (0.036)	-0.061* (0.033)	-0.092** (0.039)	-0.109** (0.053)	0.015 (0.092)
Had diarrhea in the past two weeks	-0.146** (0.072)	-0.014 (0.079)	0.000 (0.061)	-0.126*** (0.038)	-0.120*** (0.044)	-0.148*** (0.041)	-0.222*** (0.047)	-0.358*** (0.061)	-0.422*** (0.085)

Table A.6 – continued from previous page

	$Q = 0.025$ (\mathcal{M}_{A35})	$Q = 0.05$ (\mathcal{M}_{A36})	$Q = 0.10$ (\mathcal{M}_{A37})	$Q = 0.25$ (\mathcal{M}_{A38})	$Q = 0.50$ (\mathcal{M}_{A39})	$Q = 0.75$ (\mathcal{M}_{A40})	$Q = 0.90$ (\mathcal{M}_{A41})	$Q = 0.95$ (\mathcal{M}_{A42})	$Q = 0.975$ (\mathcal{M}_{A43})
	(0.069)	(0.062)	(0.052)	(0.030)	(0.033)	(0.039)	(0.040)	(0.048)	(0.080)
C. Parents' characteristics (X_P)									
Mother's BMI	0.037*** (0.007)	0.036*** (0.005)	0.033*** (0.006)	0.041*** (0.004)	0.036*** (0.004)	0.042*** (0.004)	0.050*** (0.006)	0.056*** (0.006)	0.053*** (0.009)
Mother's height	0.029*** (0.004)	0.040*** (0.003)	0.046*** (0.003)	0.049*** (0.002)	0.050*** (0.002)	0.053*** (0.003)	0.051*** (0.003)	0.047*** (0.004)	0.039*** (0.005)
Nr. years of mother's education	0.030*** (0.006)	0.032*** (0.005)	0.035*** (0.005)	0.023*** (0.003)	0.013*** (0.004)	0.009** (0.004)	0.004 (0.004)	0.006 (0.006)	0.005 (0.007)
Mother is currently working	0.109** (0.046)	0.054 (0.045)	0.016 (0.037)	-0.058** (0.027)	-0.067** (0.032)	-0.081** (0.035)	-0.063 (0.052)	-0.077 (0.054)	-0.117* (0.065)
Mother took iron during pregnancy	0.139*** (0.044)	0.101*** (0.039)	0.102** (0.041)	0.090*** (0.029)	0.054 (0.035)	-0.027 (0.035)	-0.065 (0.043)	-0.142** (0.058)	-0.147** (0.064)
Mother is smoking	0.028 (0.054)	-0.024 (0.051)	-0.034 (0.046)	-0.055 (0.043)	-0.026 (0.043)	-0.114** (0.046)	-0.126*** (0.048)	-0.255*** (0.083)	-0.098 (0.080)
State fixed effects	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Nr. observations	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659	14,659

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